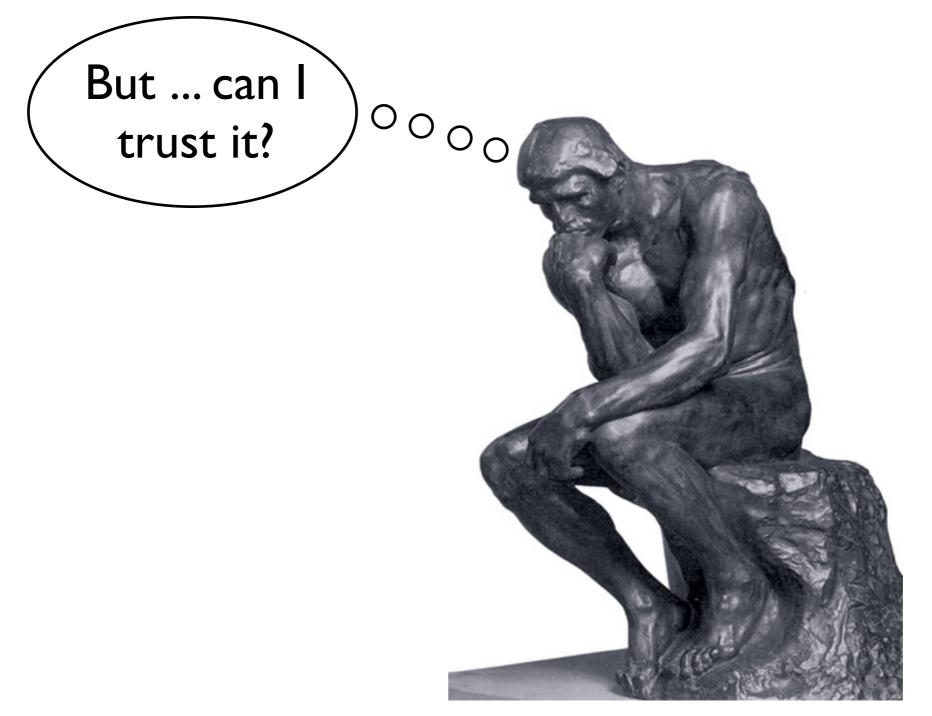


### Inference

how surprising is your statistic? (thresholding)





### Outline

- Null-hypothesis and Null-distribution
- Multiple comparisons and Family-wise error
- Different ways of being surprised
  - Voxel-wise inference (Maximum z)
  - Cluster-wise inference (Maximum size)
- Parametric vs non-parametric tests
- Enhanced clusters
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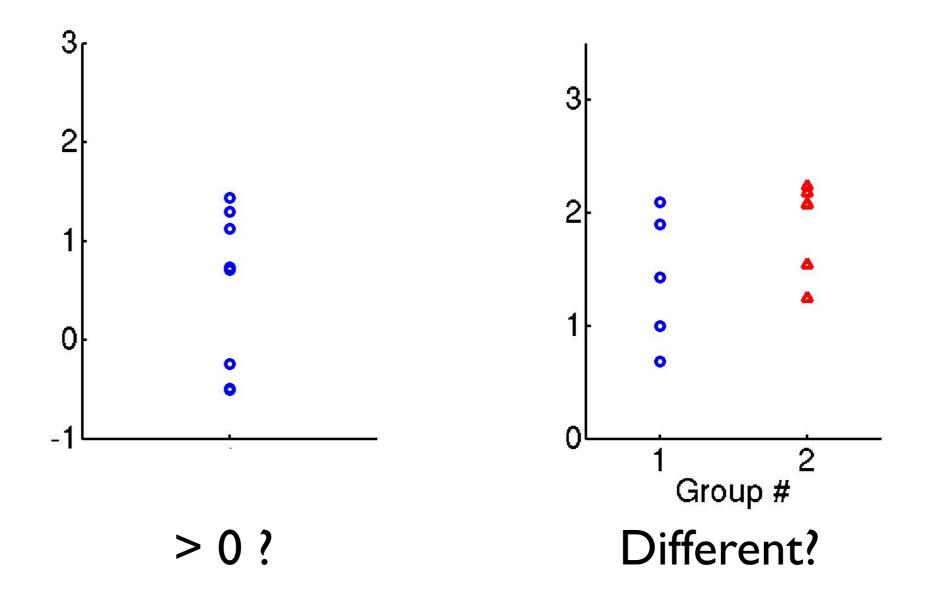
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### The task of classical inference

• Given some data we want to know if (e.g.) a mean is different from zero or if two means are different



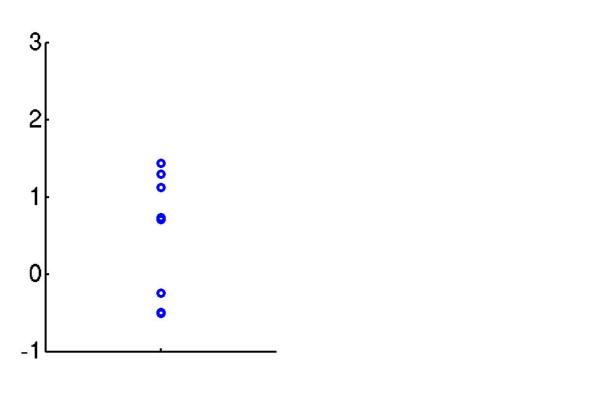


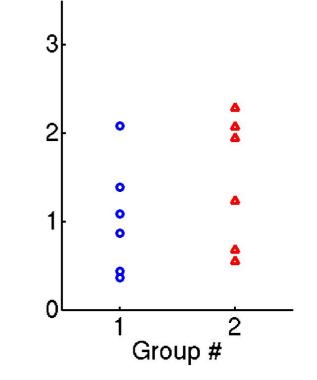
I.A null-hypothesis

Typically the opposite of what we actually "hope", e.g.

There is **no** effect of treatment:  $\mu = 0$ 

There is **no** difference between groups:  $\mu_1 = \mu_2$ 

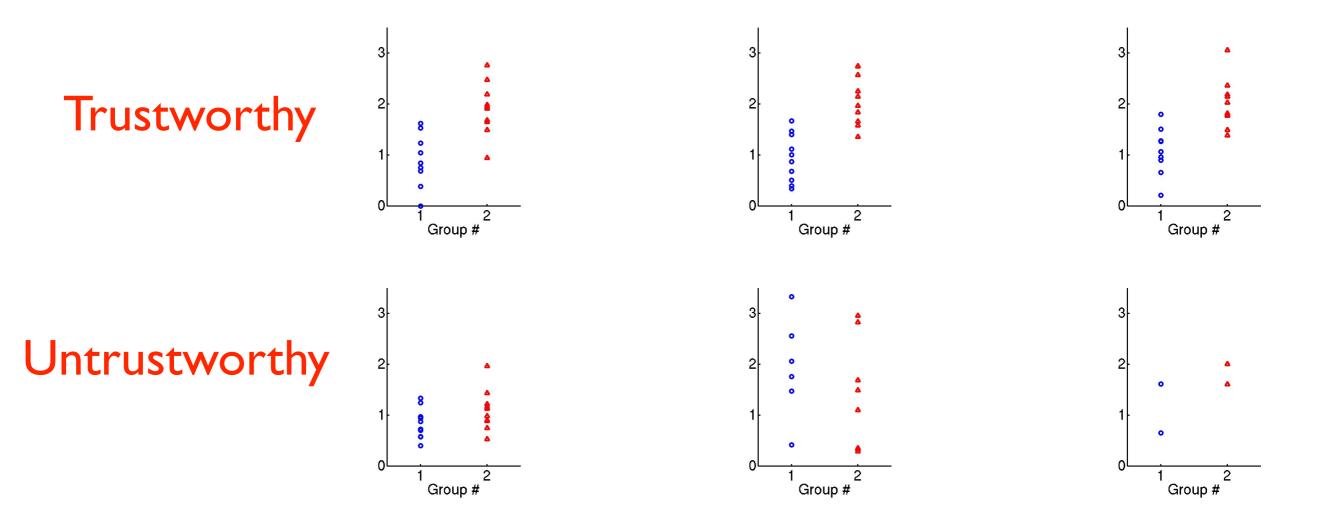






- I.A null-hypothesis
- 2. A test-statistic

Assesses "trustworthiness"





I.A null-hypothesis

t =

2. A test-statistic

Assesses "trustworthiness"

A *t*-statistic reflects precisely this

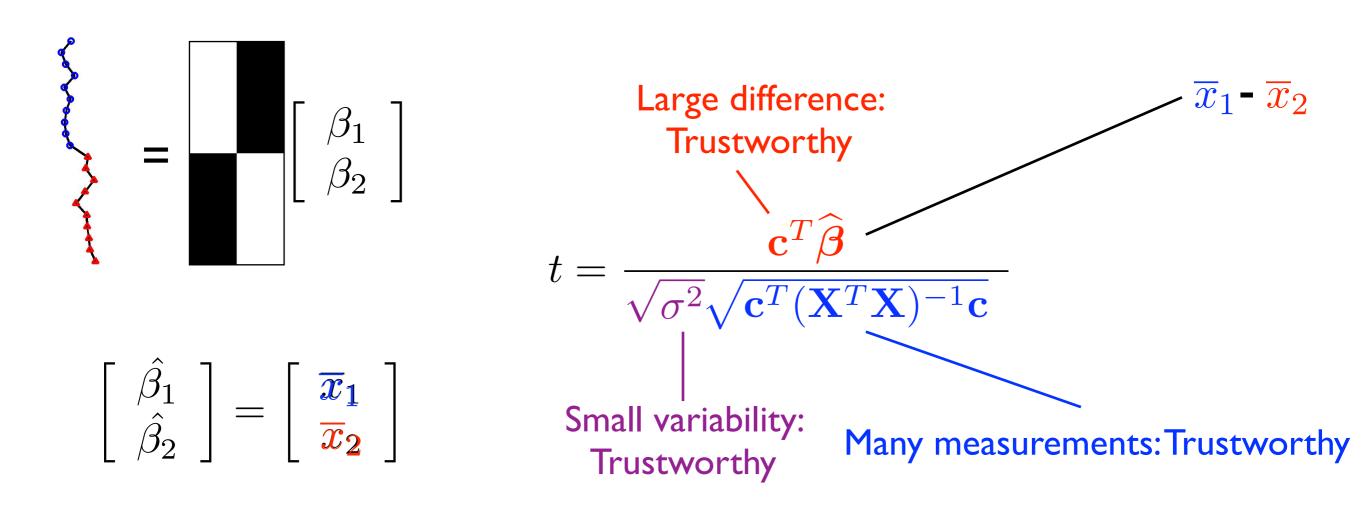
Large difference: Trustworthy

Many measurements: Trustworthy Small variability: Trustworthy



- I.A null-hypothesis
- 2. A test-statistic

Or expressed in GLM lingo





- I.A null-hypothesis
- 2. A test-statistic
- 3. A null-distribution

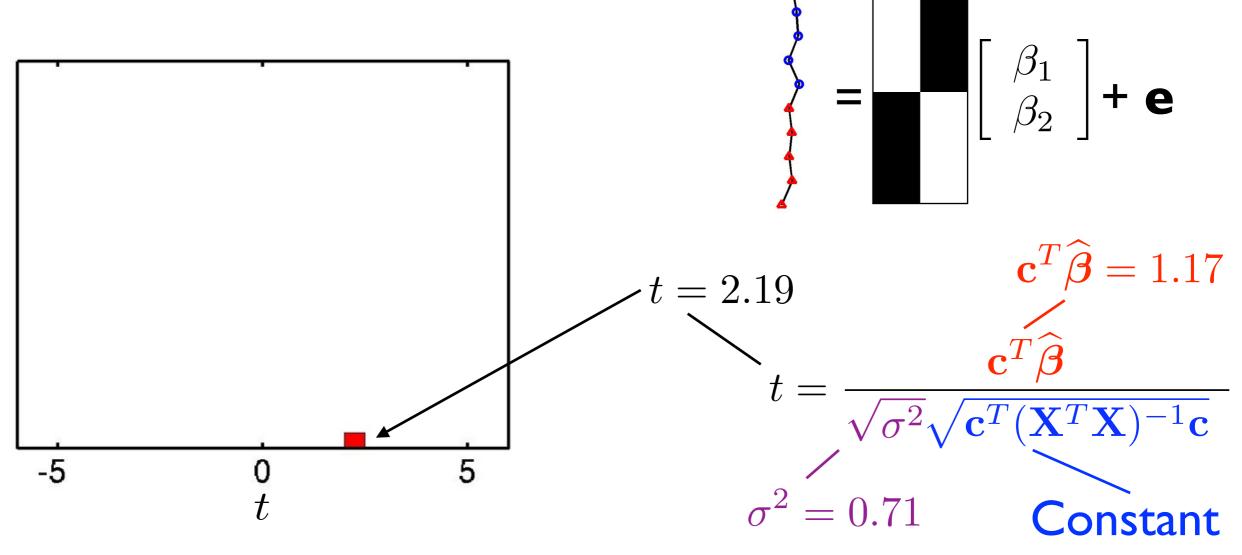
Let us assume there is no difference, i.e. the null-hypothesis is true.

We might then get these data  $\begin{bmatrix} \beta_1 \\ \beta_2 \end{bmatrix} + \mathbf{e}$ = 1.17t = 2.19



- I.A null-hypothesis
- 2. A test-statistic
- 3. A null-distribution

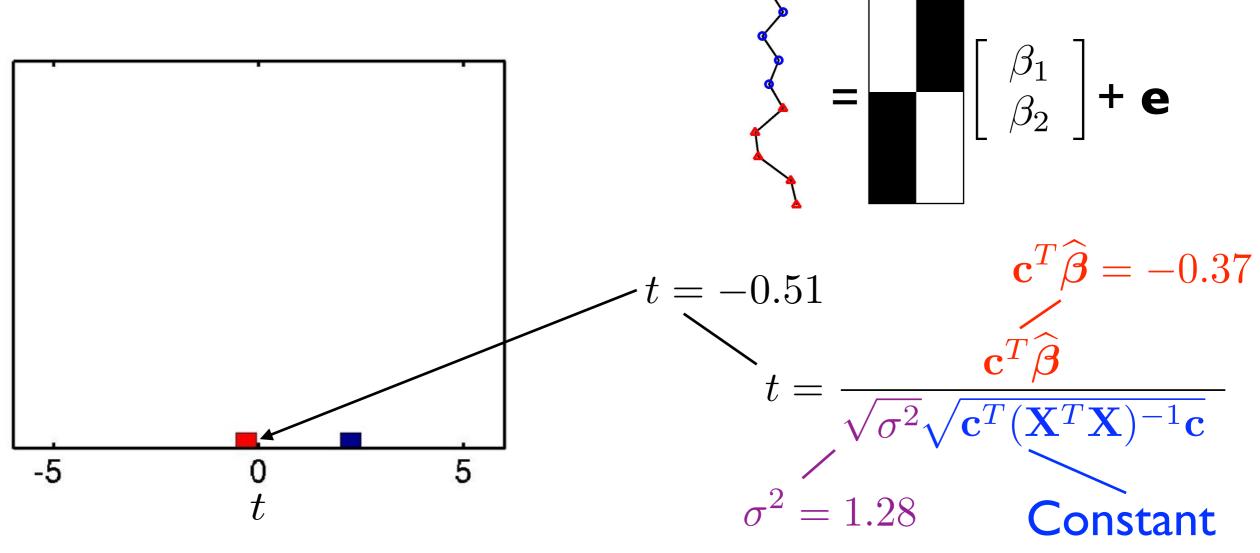
We might then get these data





- I.A null-hypothesis
- 2. A test-statistic
- 3. A null-distribution

or we could have gotten these





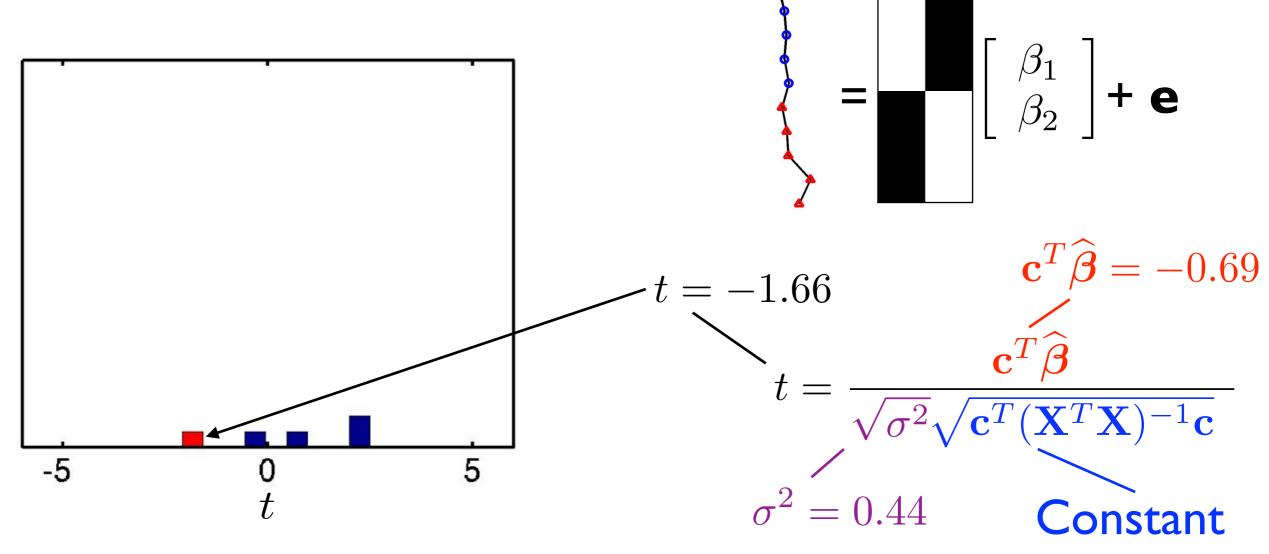
I.A null-hypothesis 2. A test-statistic maybe these 3. A null-distribution  $\begin{bmatrix} \beta_1 \\ \beta_2 \end{bmatrix} + \mathbf{e}$ = 0.31t = 0.49 $\sigma^2$ -5 5 0 onstant



I.A null-hypothesis 2. A test-statistic or perhaps these 3. A null-distribution  $\begin{vmatrix} \beta_1 \\ \beta_2 \end{vmatrix} + \mathbf{e}$ = = 1.22t = 2.19 $\sigma^2$ -5 5 0 onstant



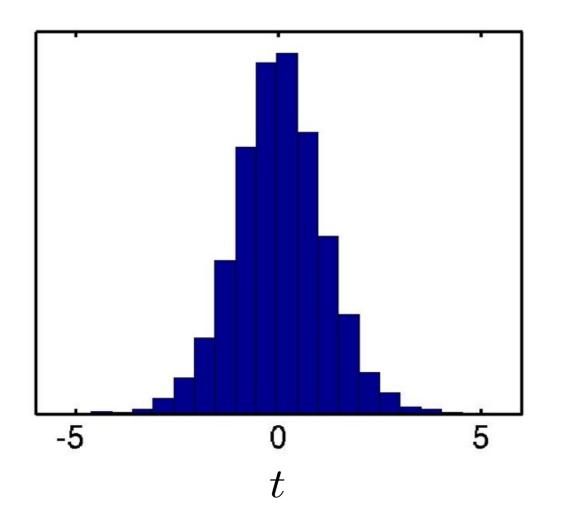
- I.A null-hypothesis
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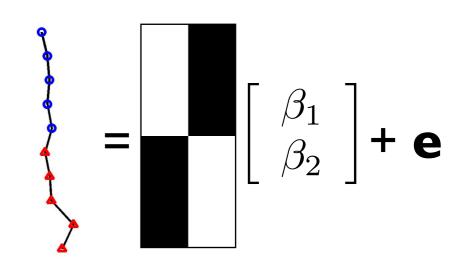


etc



- I.A null-hypothesis
- 2. A test-statistic
- 3. A null-distribution

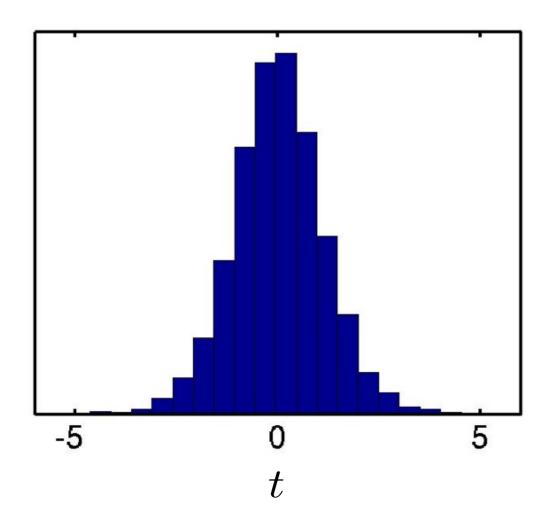




And if we do this many many many many times...



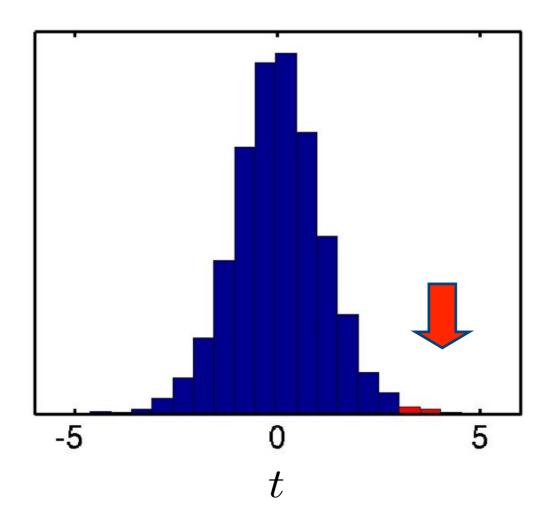
- I.A null-hypothesis
- 2. A test-statistic
- 3. A null-distribution



So, why is this helpful?



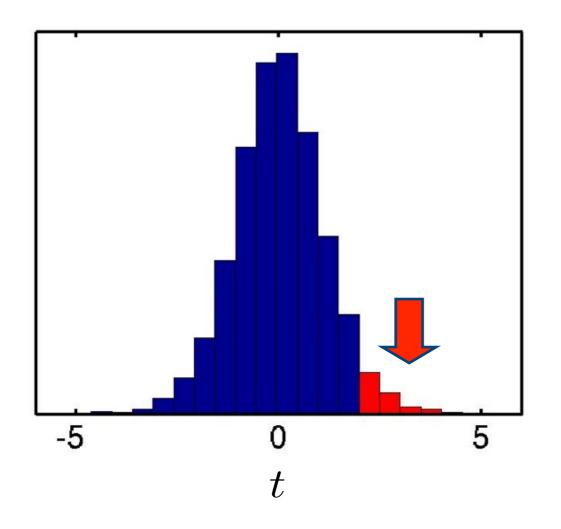
- I.A null-hypothesis
- 2. A test-statistic
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Well, it for example tells us that in ~1% of the cases t > 3.00, even when the null-hypothesis is true.



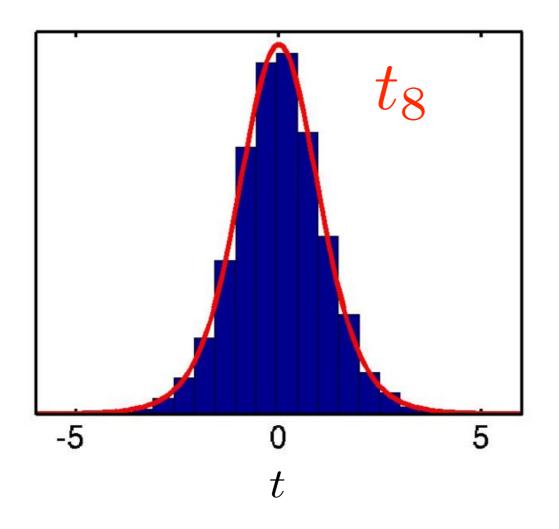
- I.A null-hypothesis
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Or that in ~5% of the cases t > 1.99. When the nullhypothesis is true.



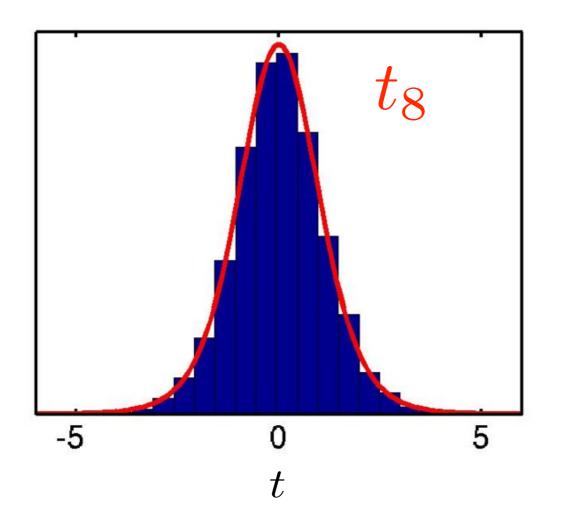
- I.A null-hypothesis
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And best of all:This distribution is known *i.e.* one can calculate it. Much as one can calculate sine or cosine



- I.A null-hypothesis
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And best of all: This distribution is known *i.*e. one can calculate it. Much as one can calculate sine or cosine

Provided that  $\mathbf{e} \sim N(0,\sigma^2)$ 

I.A null-hypothesis

$$H_0: \overline{x}_1 = \overline{x}_2$$
,  $H_1: \overline{x}_1 > \overline{x}_2$ 

- 2. A test-statistic
- 3. A null-distribution

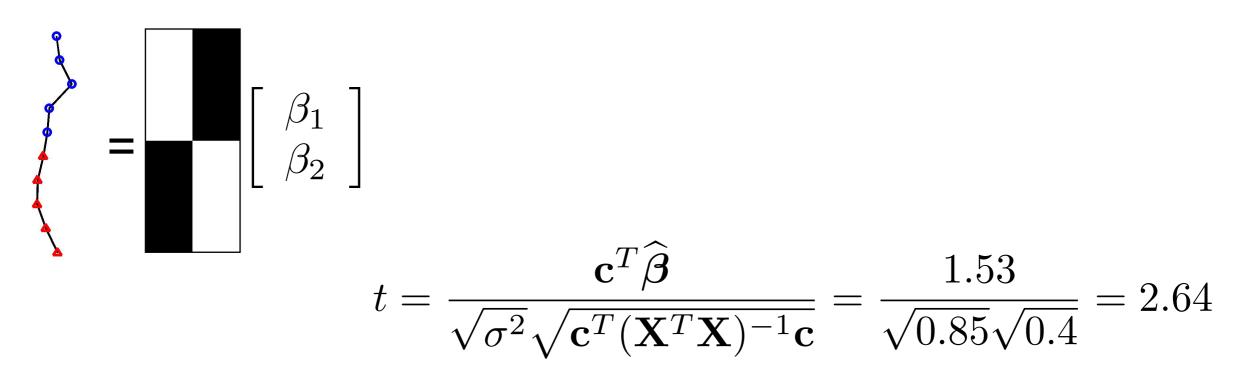
So, with these tools let us do an experiment

 $H_0: \overline{x}_1 = \overline{x}_2$ ,  $H_1: \overline{x}_1 > \overline{x}_2$ 

 $t_8 = 2.64$ 

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- 2. A test-statistic
- 3. A null-distribution

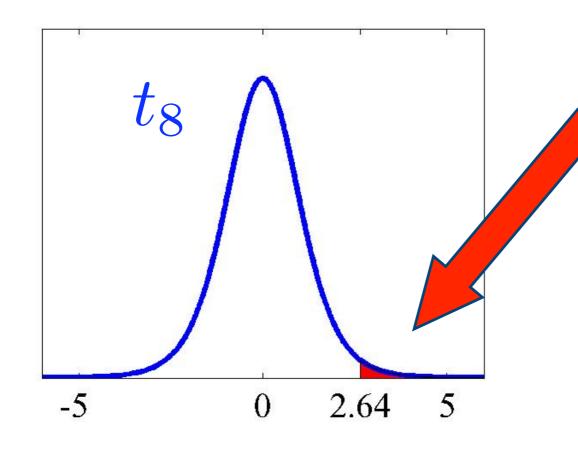
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So, with these tools let us do an experiment

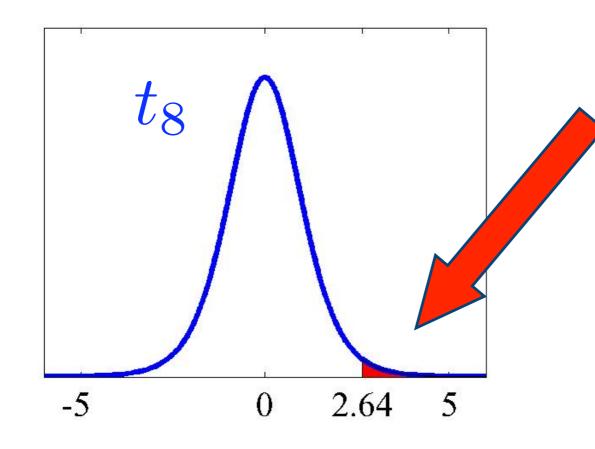


If the null-hypothesis is true, we would expect to have a ~1.46% chance of finding a t-value this large or larger

- I.A null-hypothesis
- 2. A test-statistic
- 3. A null-distribution

$$H_0: \overline{x}_1 = \overline{x}_2 , H_1: \overline{x}_1 > \overline{x}_2$$
$$t_8 = 2.64$$
$$t_8 = 2.64*$$

So, with these tools let us do an experiment



There is ~1.46% risk that we reject the nullhypothesis (i.e. claim we found something) when the null is actually true. We can live with that.

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- But what does that actually mean?

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- We want to perform an experiment and as part of that we define a null-hypothesis, e.g.  $H_0: \mu = 0$
- Now what can happen?

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 $H_0$  is true  $H_0$  is false  $H_0$  is false

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 $H_0$  is true B True state of affairs  $H_0$  is false B

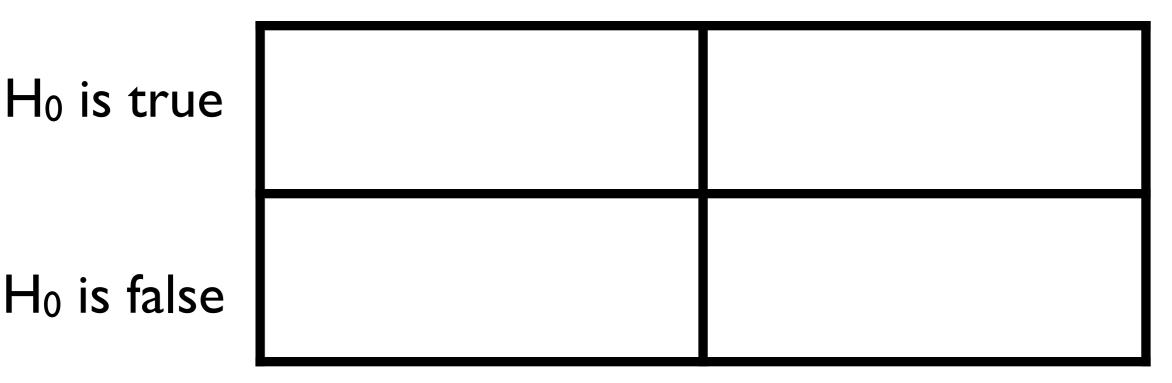
We don't reject  $H_0$  We reject  $H_0$  Our decision



 $H_0$  is true B True state of affairs  $H_0$  is false B

We don't reject  $H_0$  Gur Gur

We don't reject  $H_0$  We reject  $H_0$ 





 $H_0$  is true  $H_0$  is false  $H_0$  is false

We don't reject  $H_0$  Gur Gur

We don't reject  $H_0$  We reject  $H_0$ 

H<sub>0</sub> is true  $\bigcirc$  H<sub>0</sub> is false



 $H_0$  is true  $H_0$  is false  $H_0$  is false

We don't reject  $H_0$  Gur decision We reject  $H_0$ 

We don't reject  $H_0$  We reject  $H_0$ 

$H_0$ is true		False positive
H <sub>0</sub> is false	False negative	



 $H_0$  is true B True state of affairs  $H_0$  is false B

We don't reject  $H_0$  Gur Gur

We don't reject  $H_0$  We reject  $H_0$ 

$H_0$ is true		False positive Type I error
H <sub>0</sub> is false	False negative Type II error	



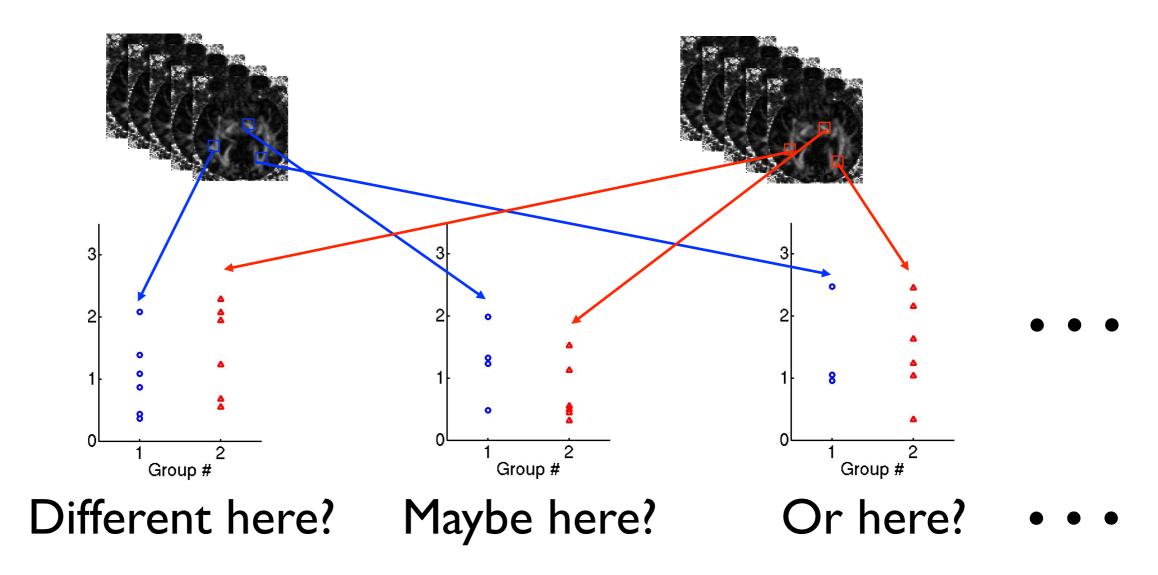
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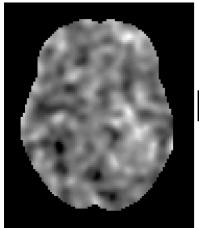
### Multiple Comparisons

 In neuroimaging we typically perform <u>many</u> tests as part of a study

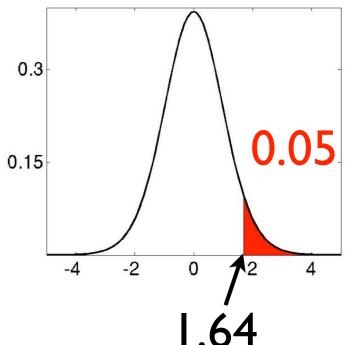




### What happens when we apply this to imaging data?

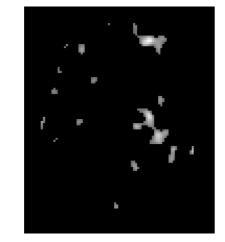


z-map where each voxel ~N. Null-hypothesis true everywhere, i.e. NO ACTIVATIONS



Ζ

z-map thresholded at I.64



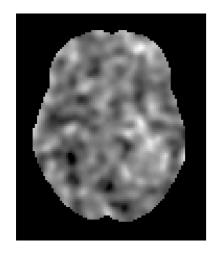
16 clusters288 voxels~5.5% of the voxels

That's a LOT of false positives

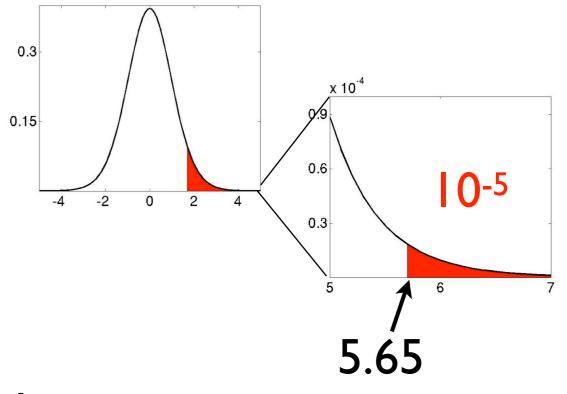


#### The strict approach: Bonferroni correction

#### Bonferroni says threshold at $\alpha$ divided by # of tests



5255 voxels 0.05/5255≈10-5

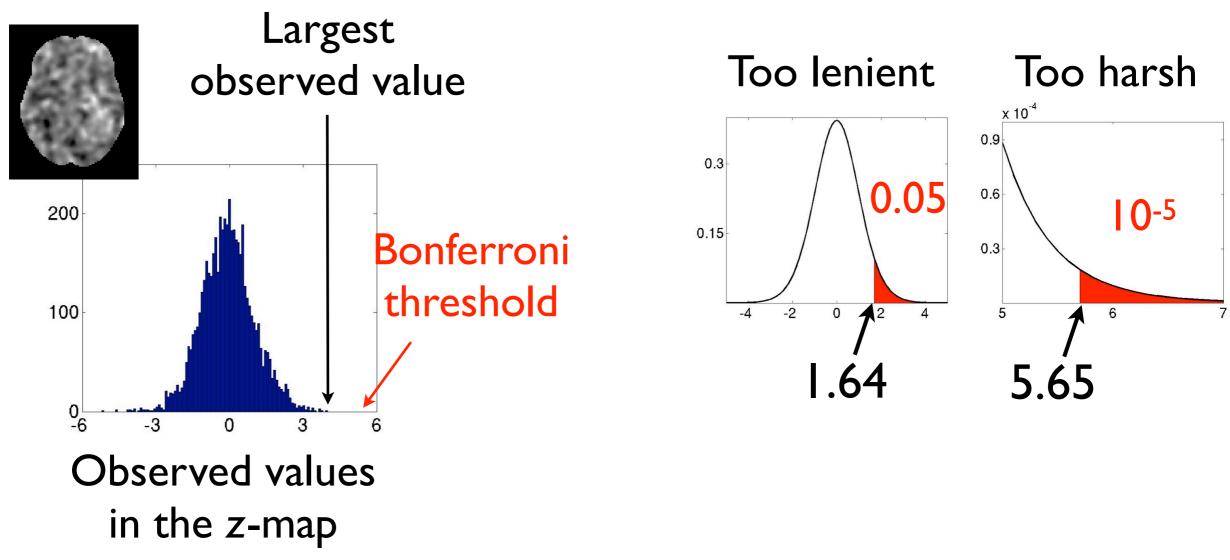


z-map thresholded at 5.65





#### But ... doesn't 5.65 sound very high?

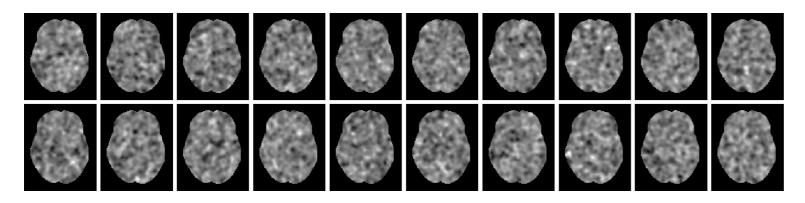


So what do we want then?



## Family-wise error

Let's say we perform a series of identical studies



Each z-map is the end result of a study

Let us further say that the null-hypothesis is true We want to threshold the data so that only once in 20 studies do we find a voxel above this threshold



But how do we find such a threshold?

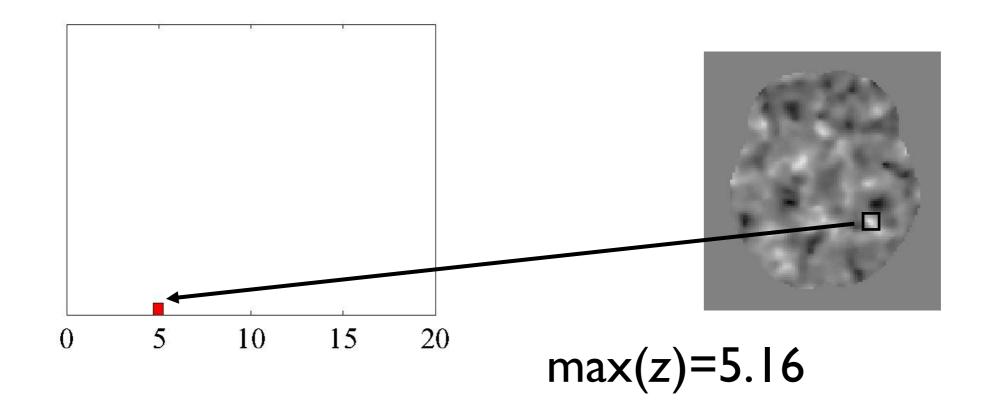


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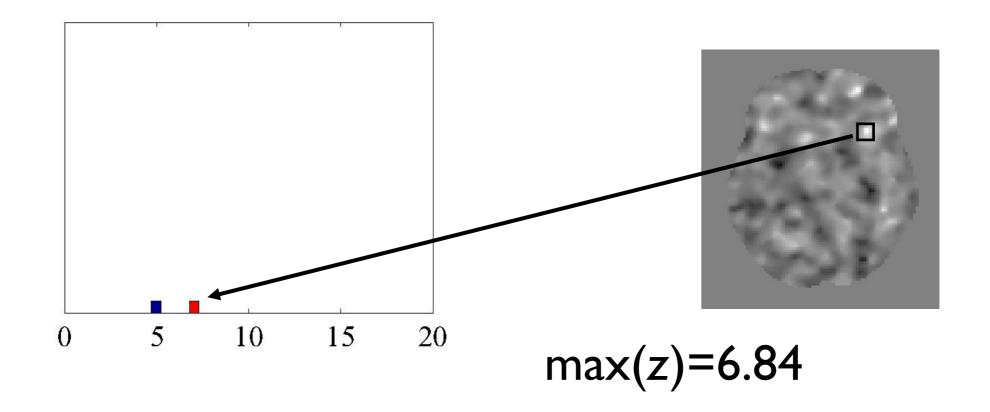


- When we want to control "family-wise error", what do we in practice want?
- If the null-hypothesis is true (no activation) we want to reject it no more than 5% of the time.
- And if we reject anything, we will definitely reject the most "extreme" value (max(z)) in the brain.



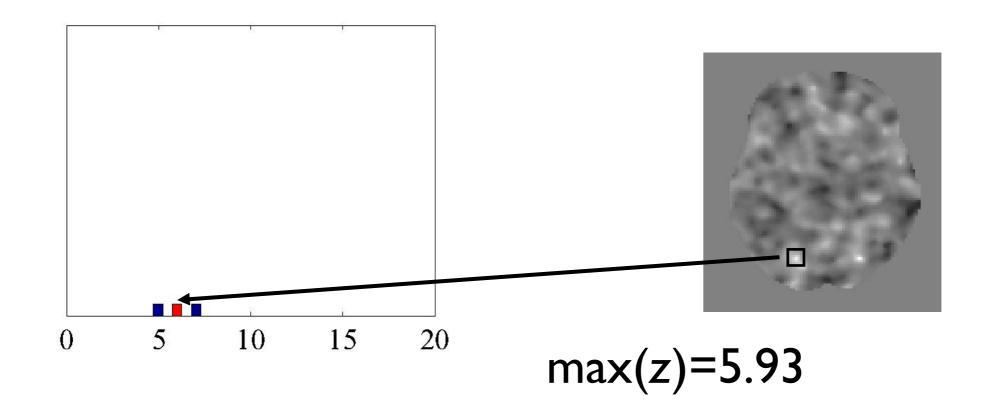


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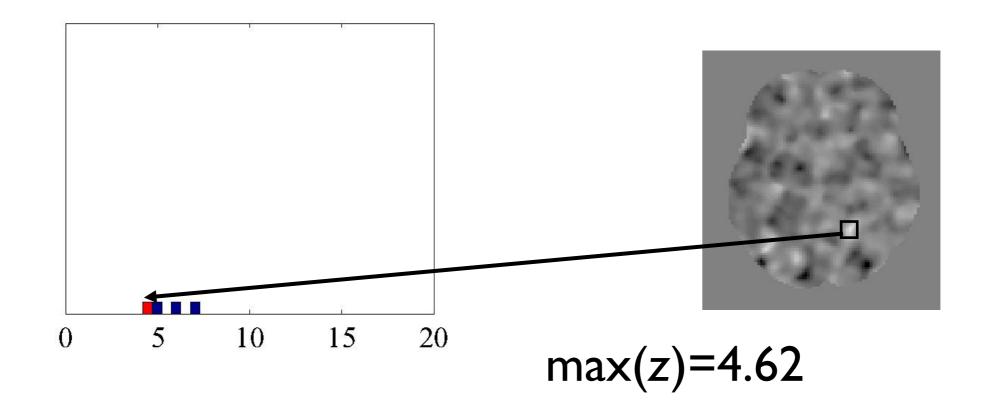


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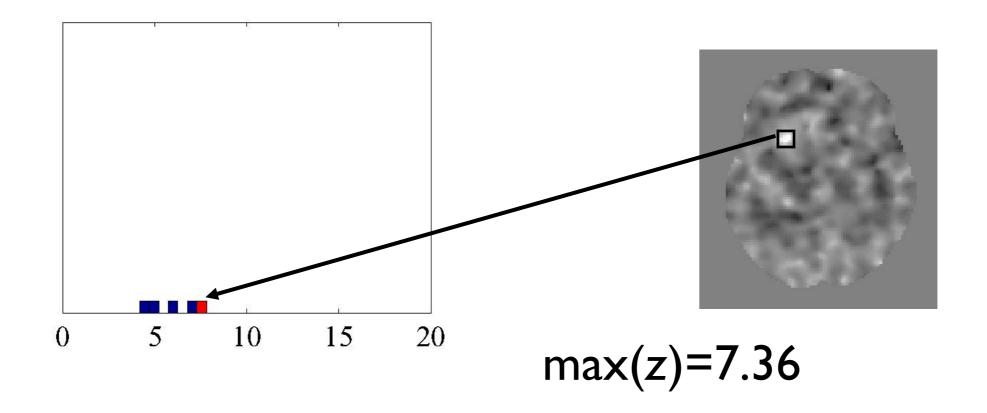


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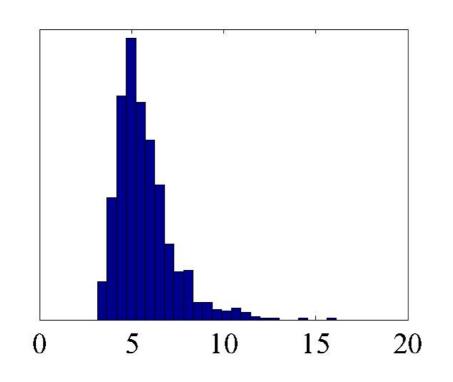
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## Maximum Z

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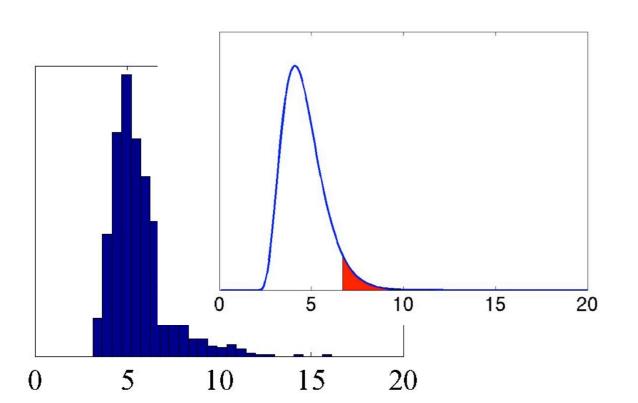






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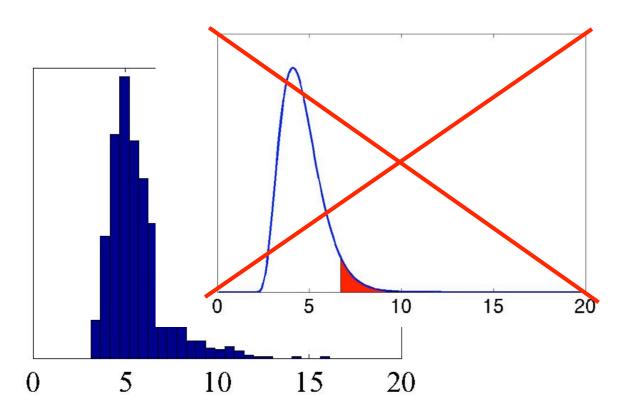


This is the distribution we want to use for our FWE control.



## Maximum Z

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This is the distribution we want to use for our FWE control. But there is no known expression for it! 🛞

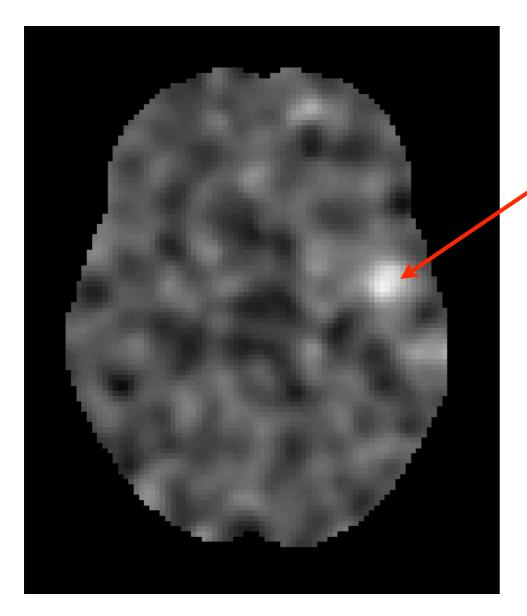


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# Spatial extent: another way to be surprised

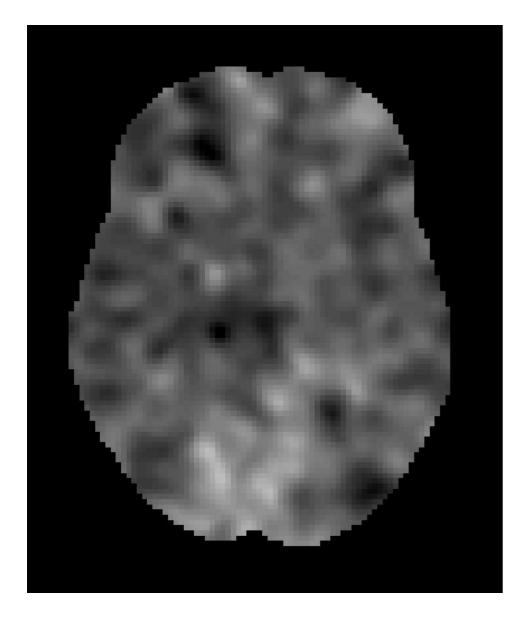
This far we have talked about voxel-based tests



We say: Look! A z-value of 7. That is so surprising (under the nullhypothesis) that I will have to reject it. (Though we are of course secretly delighted to do so)

# Spatial extent: another way to be surprised

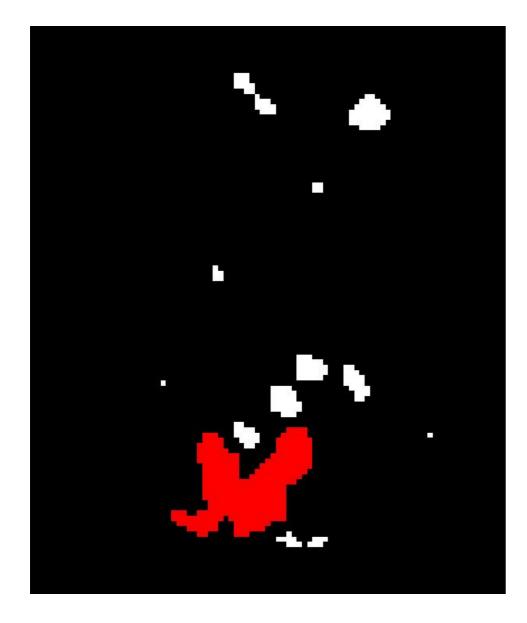
But sometimes our data just aren't that surprising.



Nothing surprising here! The largest z-value is ~4. We cannot reject the null-hypothesis, and we are **devastated**.

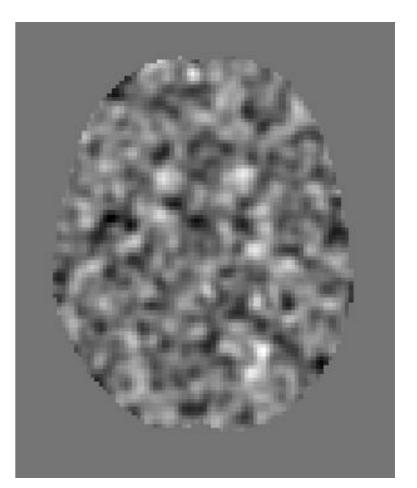
# Spatial extent: another way to be surprised

So we threshold the z-map at 2.3 (arbitrary threshold) and look at the spatial extent of clusters



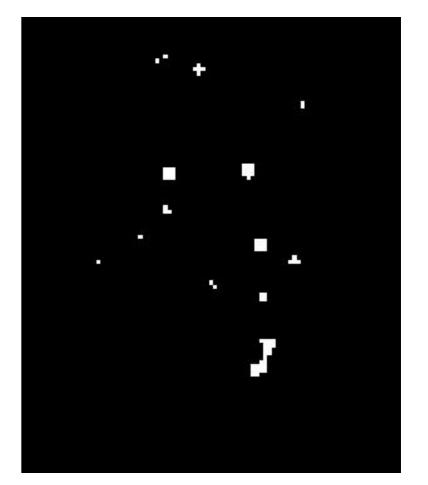
We say: Look at that whopper! 301 connected voxels all with z-values > 2.3.That is really surprising (under the null-hypothesis). I will have to reject it.

As with the *z*-values we need a "null-distribution". What would that look like in this case?



Let's say we have acquired some data

If we reject any cluster we will reject the largest. So what we want is the distribution of the largest cluster, under the null-hypothesis.



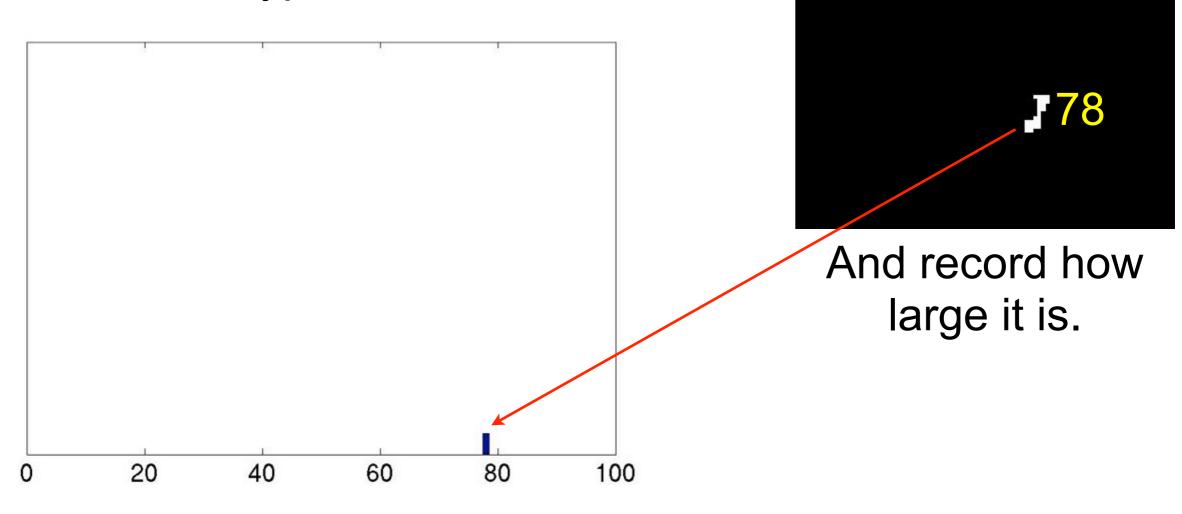
Threshold the z-map at 2.3 (arbitrary)

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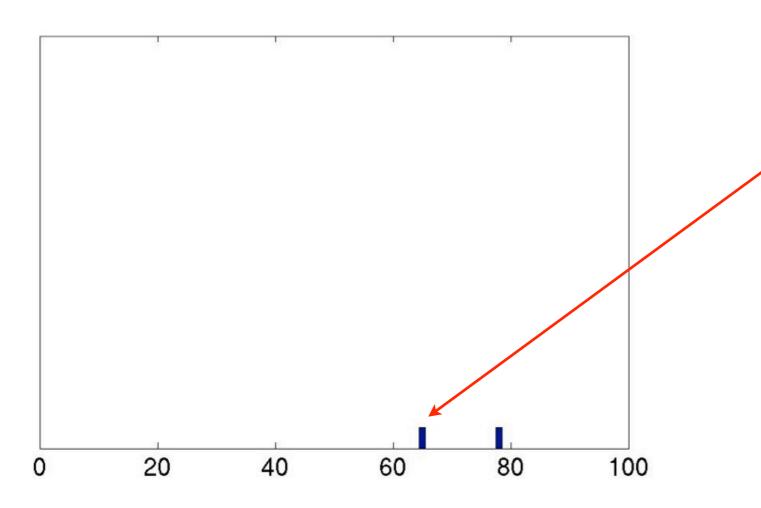


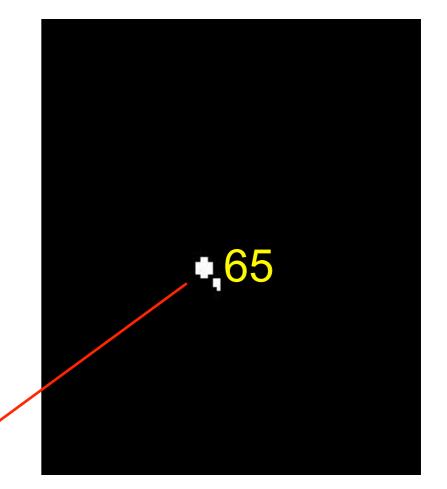
Locate the largest cluster anywhere in the brain.

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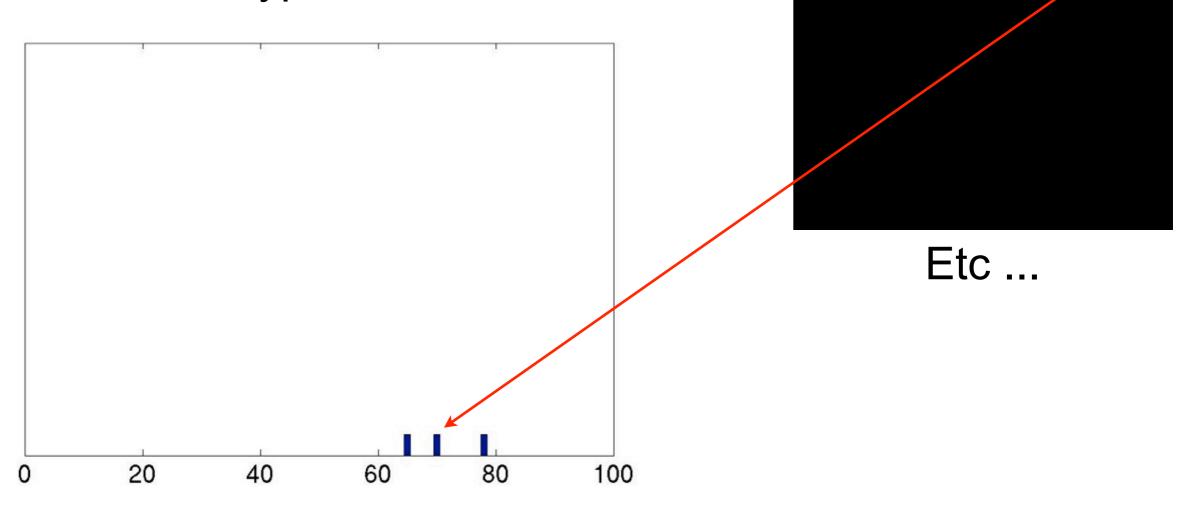
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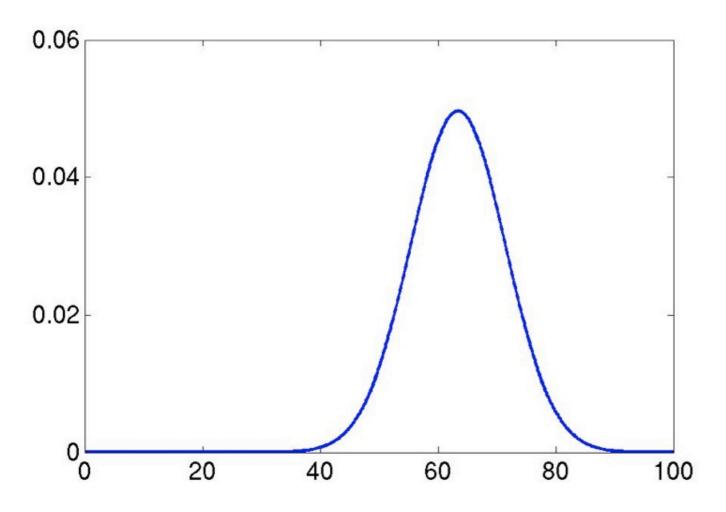


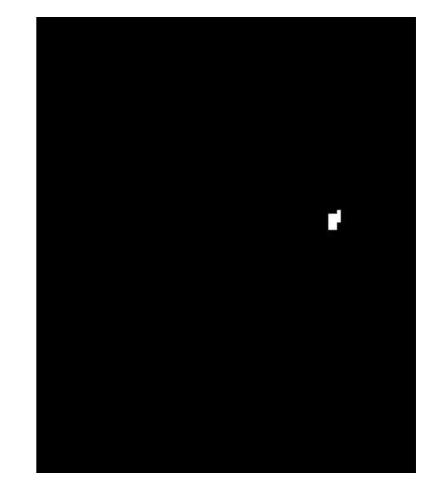
And do the same for another experiment...

If we reject any cluster we will reject the largest. So what we want is the distribution of the largest cluster, under the null-hypothesis.



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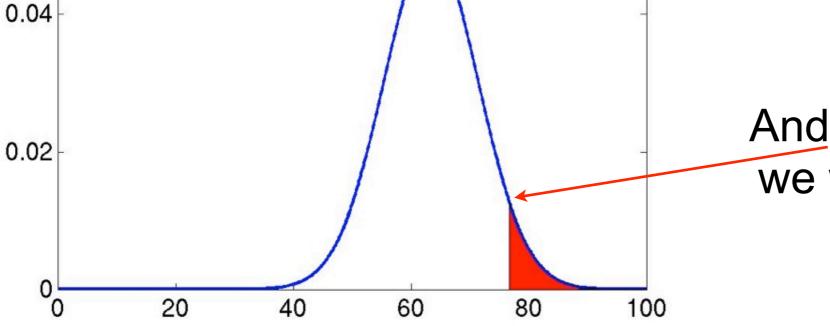


Until we have ...

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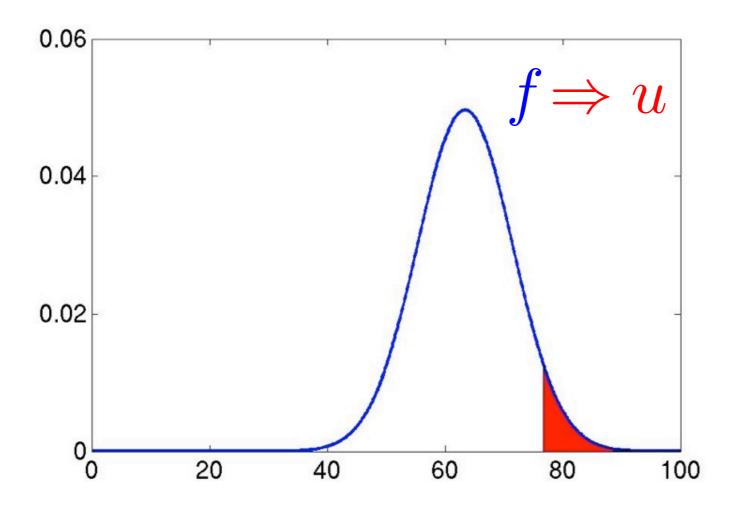
0.06

If we find a cluster larger than 76 voxels we reject the null-hypothesis.



And this (76) is the level we want to threshold at

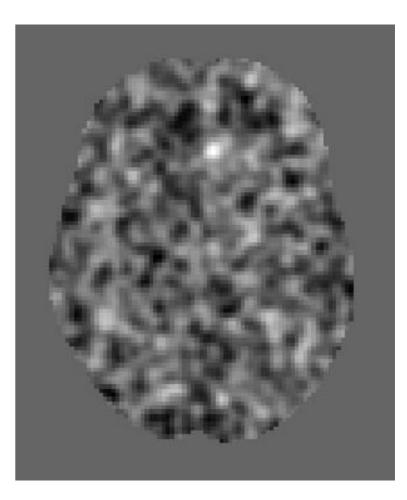
So, just as was the case for the tvalues, we now have a distribution *f* that allows us to calculate a Family Wise threshold *u* pertaining to cluster size.



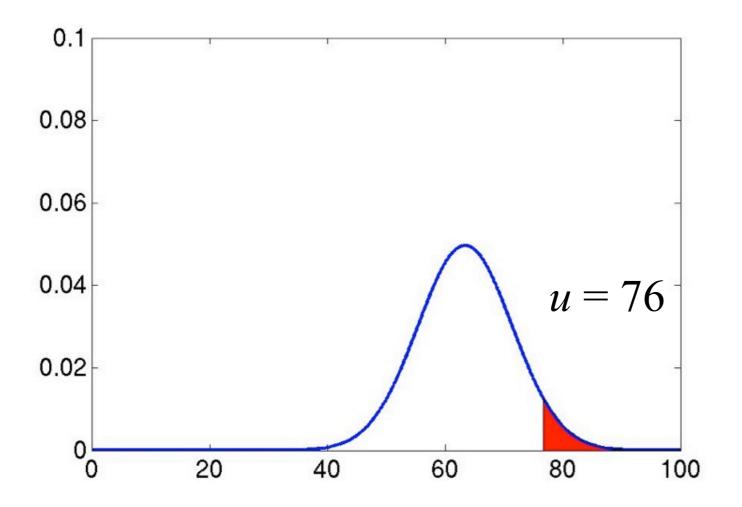
But what does f and u crucially depend on?

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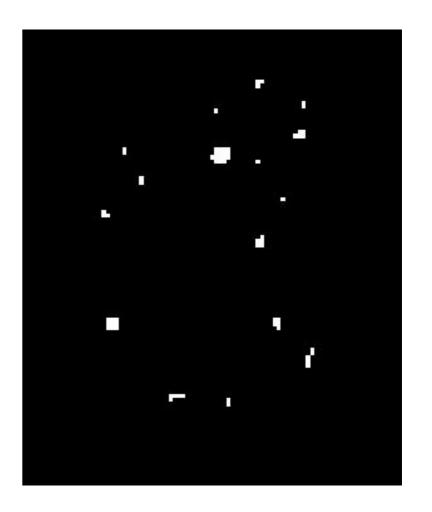
f depends crucially on the initial "clusterforming" threshold?



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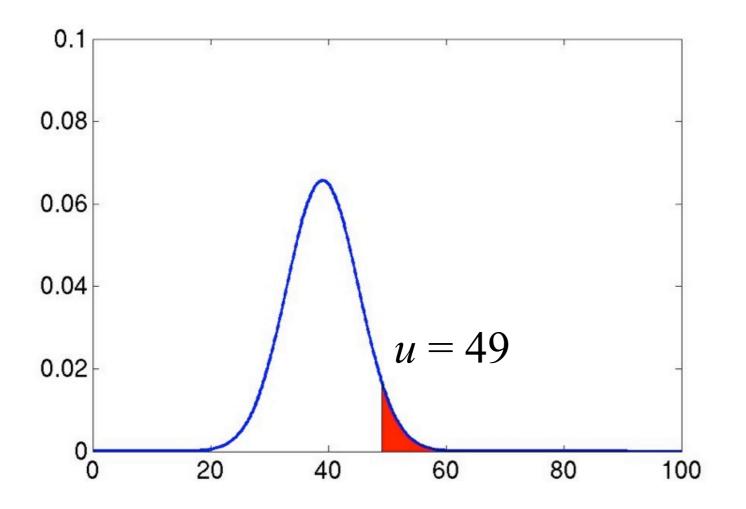


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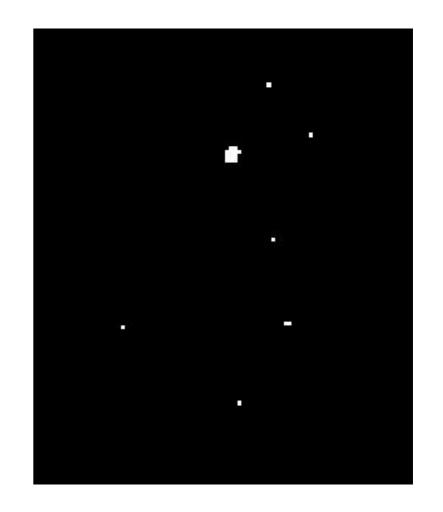


*z* = 2.3

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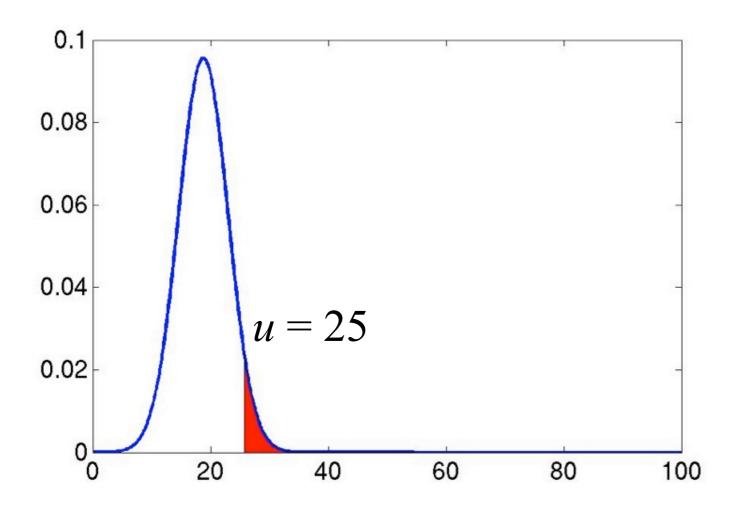


f depends crucially on
 the initial "cluster forming" threshold?

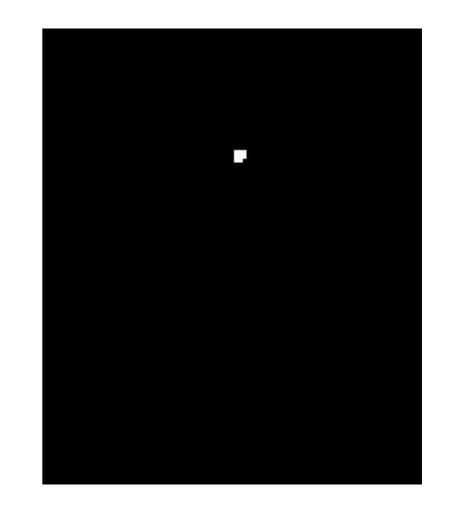


z = 2.7

So, just as was the case for the zvalues, we now have a distribution *f* that allows us to calculate a Family Wise threshold *u* pertaining to cluster size.

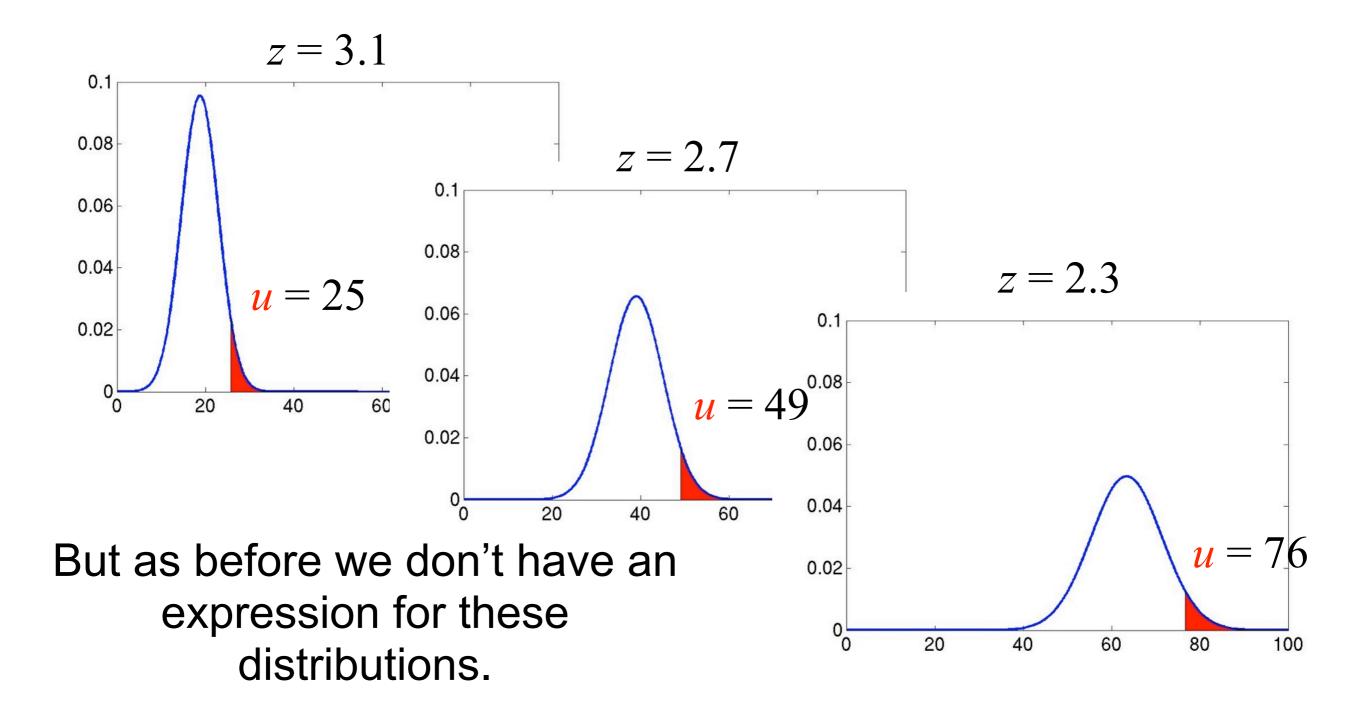


f depends crucially on
 the initial "cluster forming" threshold?



z = 3.1

Hence the distribution for the cluster size should really be written f(z) and the same for u(z)



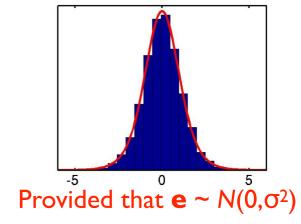


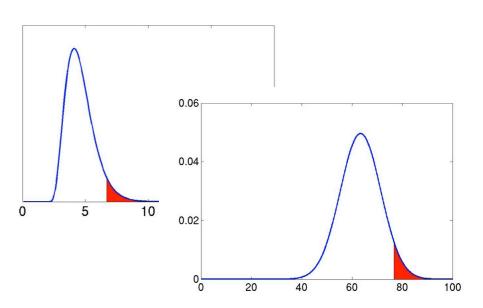
## Outline

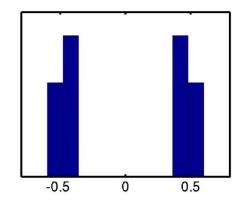
- Null-hypothesis and Null-distribution
- Multiple comparisons and Family-wise error
- Different ways of being surprised
  - Voxel-wise inference (Maximum z)
  - Cluster-wise inference (Maximum size)
- Parametric vs non-parametric tests
- Enhanced clusters
- FDR False Discovery Rate

- As we described earlier, one of the great things about for example the t-test is that we know the nulldistribution
- But most distributions are not that simple

 And errors are not always normaldistributed

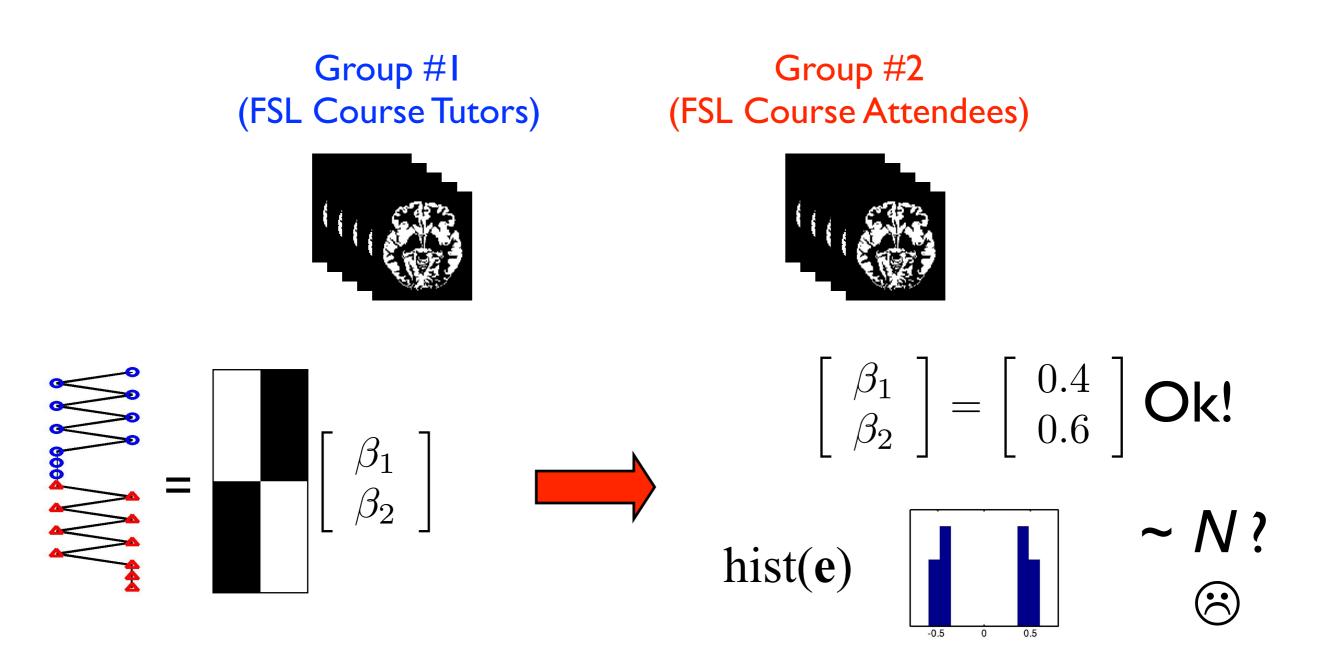






## Example: VBM-style analysis

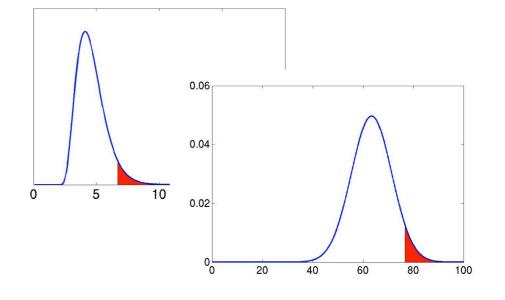
- Our data is segmented grey matter maps
- A voxel is either grey matter, or not.



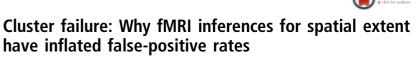
• There are <u>approximations</u> to the Max-z and Max-size statistics

 These are valid under certain sets of assumptions

But can be a problem when applied outside of that set of assumptions



- Search area "large relative to boundary"
- "High enough" cluster forming threshold
- Normal distributed errors



Anders Eklund<sup>a,b,c,1</sup>, Thomas E. Nichols<sup>d,e</sup>, and Hans Knutsson<sup>a,c</sup>

<sup>a</sup>Division of Medical Informatics, Department of Biomedical Engineering, Linköping University, 5-581 85 Linköping, Sweden; <sup>b</sup>Division of Statistics and Machine Learning, Department of Computer and Information Science, Linköping University, 5-581 83 Linköping, Sweden; <sup>c</sup>Center for Medical Image Science and Visualization, Linköping University, 5-581 83 Linköping, Sweden; <sup>c</sup>Center for Medical Image Kingdom; and <sup>®</sup>WMG, University of Warwick, Coventry CV4 7AL, United Kingdom

Edited by Emery N. Brown, Massachusetts General Hospital, Boston, MA, and approved May 17, 2016 (received for review February 12, 2016)

The most widely used task functional magnetic resonance imagin (fMRI) analyses use parametric statistical methods that depend on variety of assumptions. In this work, we use real resting-state dat and a total of 3 million random task group analyses to comput empirical familywise error rates for the fMRI software packages SPM WE), the chance of one or more false positives, and empirically easure the FWE as the proportion of analyses that give rise to y significant results. Here, we consider both two-sample and e-sample designs. Because two groups of subjects are randomly awn from a large group of healthy controls, the null hypothesis

 Those approximations were based on Gaussian Random Field Theory, and was an impressive body of work

• They served us fantastically well at a time when we had little choice

 But the we've moved towards nonparametric testing

Averal of Basician Excit 12, 54, 546-409 ESTIMATING THE NUMBER OF FIELD USING THE HADWIGER EXCURSION SETS, WITH A MEDICAL IMA	LOCA FIELD PEAKS IN A RANDOM CHARACTERISTIC OF PPLICATIONS TO KE	NG FOR SIGNALS WITH UNKNOWN TION AND SCALE IN A <u>x</u> <sup>2</sup> RANDOM , WITH AN APPLICATION TO FMRI TH J. WORSLEY, <sup>*</sup> McGill University	
By K. J. Work	SLEY	Abstract	
McGill U Certain three-dimensional ima physics are modelled as a smooth 1 interested in the number of peaks or This paper studies the Hadwiger ch random field. The excursion set is exceeds a fixed threshold, and the H characteristic, counts the number of sion set minus the number of "holes	The Geometry	of Random Images	nals with unknown st statistic was the ice', N dimensions ace is identical to a ough the emphasis ient. Two methods
characteristic is a measure of the excursion sets has been studied by geometry: characteristic of excursion of the number of "upcrossings" of process. The IG characteristic equals sion set provided that the set does n and Adler found the expected IG cl		J. Worsley	= 3: one based on er characteristic of v the latter method result to $\chi^2$ fields. case. In this paper
field inside a fixed volume. Worsky, IG characteristic as an estimator of positron emission tomography (PWT) Unfortunatory the IG characteristic investigation of the state of the state investigation of the state of the state ortical regions mare the boundary of Hadwiger characteristic, which is id under totation and does count out boundary or ast. Our main result is Hadwiger characteristic for an isot and three dimensions and on a sume recommission are applied to PET	and angles but the modern geometry of topology shape. What has this to do with statistics? Some ent work has found some fascinating applications mixture of geometry, topology, probability, and at tics to some very pressing problems in newly emer areas of medicial imaging and astrophysics. Where is the link? Let us begin with a quick in duction to one of the fundamental tools of topo the Euler characteristic.	and longer holds. In fact, the result is $V - E + F - P = 0$ re for any solid with just one hole. of a Too badl But this does not deter a good this mathematican-far from $i - topoens up vast new pos- nging sibilities! What happens if there are two holes, like a figure 8 (see Fig. 4)? Then it turns out that tro: V - E + F - P = -1, and so on; each hole reducesog, V - E + F - P by 1.So now suddenly we have a fascinating new tool.We can count the number of holes in a solid using the$	in images obtained ry; image analysis;
<ol> <li>Introduction. Many studies o mography (PET) involve the inter sually the difference between two good flow under baseline and stimul</li> </ol>	Topology: The Euler Characteristic Named after Leonhard Euler (1707–1783), the i prolific mathematician of the 18th century, the ler characteristic itself began with Euler's observa about polyhedra. Recall that a polyhedron is a solid object boun	Eu- Thus is born the field of topology: We define the Euler tion characteristic (EC) of a solid as simply $\mathrm{EC}=V-E+F-P$	sion tomography (PET interested in detectin the signal to noise ration though a filter $f$ . The
Received June 1990; revised June 1994. <sup>1</sup> Besearch supported by the Natural Scieno d the Fonds pour la Formation des Chercher AMS 1991 subject classifications. Primary 6 Key words and plerases. Euler characteristis lity, image analysis. 6	by plane faces, such as a cube. Euler realized the you count the faces $(F)$ , edges $(E)$ , and vertices of a polyhedron, then $V - E + F = 2$ no matter the polyhedron is constructed. A cube, for example, has $F = 6$ faces, $E = 12$ e and $V = 8$ vertices (see Fig. 1a) so that $8 - 12 + 6$ For a solid that consists of $P$ polyhedra, stuck toge	(iv) Thus the EC of a pretzel-shaped solid (Fig. 1e) is how −2: +1 for the solid part (the part you eat), and −3 for each of the three holes, giving −2 overall. Have leges we covered all possibilities? Not quite—if the solid is = 2. hollow, like a tennis ball, then surprisingly enough the ther EC is 2 (see Fig. 1f).	(1.1 from the matched filte
	on at least one common face, the slightly more ger formula becomes $V - E + F - P = 1$ . A little experimentation will convince you that	eral Think you've got it now? How about a solid shaped like a bicycle inner tube? Answer: The EC is 0, and if	

In task on will convince you that this it has a puncture, then the  $D \subset S - I$ . is for all solids (see Fig. 1b)—well One more slight generalization, which will prove to olid has a hole going right through be extremely useful for practical applications: Suppose







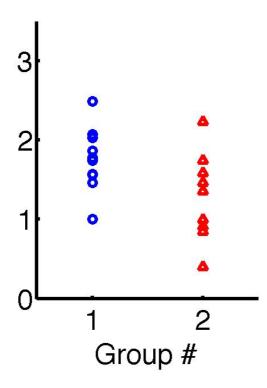


	EAT - FMRI Expert Analysis Tool v6.00	
Higher-level analysis - Statistics - Misc Data Prestats Registration Stats Post-stats Mixed effects: FLAME 1 - Use automatic outlier de-weighting Model setup wizard	Higher-level analysis - Statistics - Misc Data Prestats Registration Stats Post-stats Randomise - Permutations 5000 - Use automatic outlier de-weighting Model setup wizard	
Full model setup       Go     Save       Load     Exit       Help     Utils	Full model setup       Go     Save       Load     Exit       Help     Utils	

## A simple permutation test

- We can permute the data itself to create a distribution that we can use to test our statistic.
  - + Makes very few assumptions about the data
  - + Works for any test statistic

We have performed an experiment



And calculated a statistic, e.g. a *t*-value

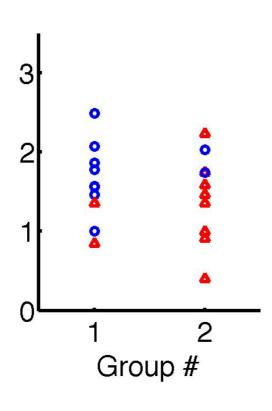
t = 2.27

If the null-hypothesis is true, there is no difference between the groups. That means we should be able to "re-label" the individual points without changing anything.

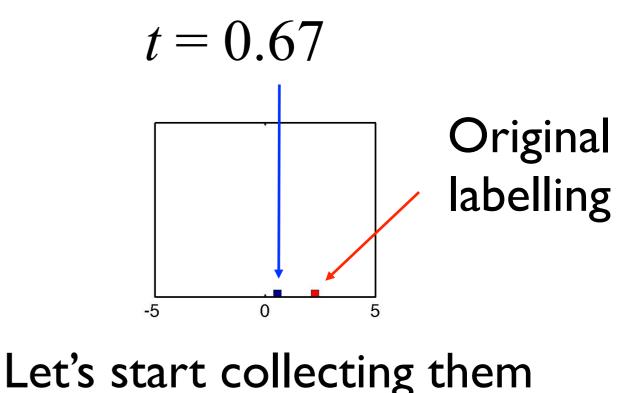
# A simple permutation test

- We can permute the data itself to create a distribution that we can use to test our statistic.
  - + Makes very few assumptions about the data
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One re-labelling





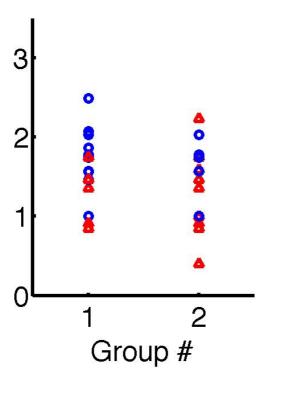


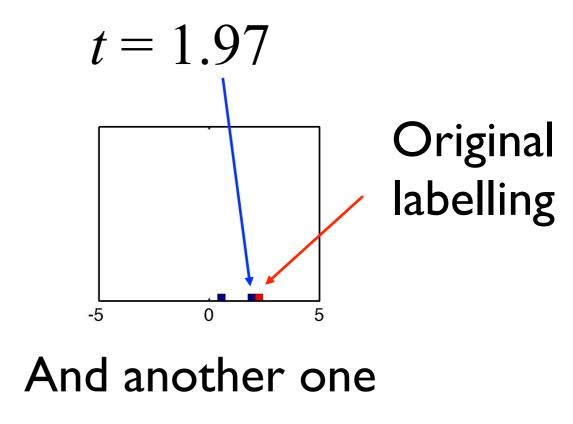
# A simple permutation test

- We can permute the data itself to create a distribution that we can use to test our statistic.
  - + Makes very few assumptions about the data
  - + Works for any test statistic

Second re-labelling

*t*-value after re-labelling





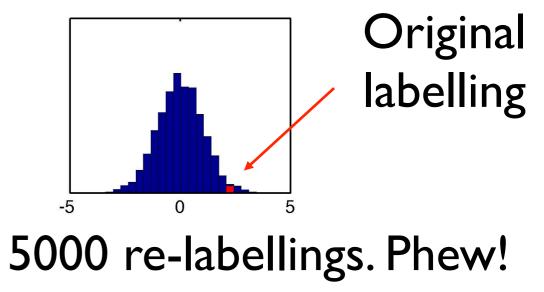
# A simple permutation test

- We can permute the data itself to create a distribution that we can use to test our statistic.
  - + Makes very few assumptions about the data
  - + Works for any test statistic

Of the 5000 re-labellings, only 90 had a tvalue > 2.27 (the original labelling).

I.e. there is only a ~1.8% (90/5000) chance of obtaining a value > 2.27 if there is no difference between the groups

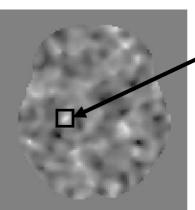
i.e. 
$$p(x \ge 2.27) = 1.79\%$$
 for  $t_{18}$ 



# And we can use this for any statistic

This is what we got

We compared activation by painful stimuli in two groups of 5 subjects each.



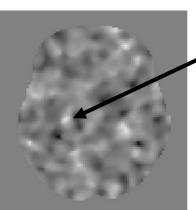
<u>Very</u> intriguing activation.  $t_8 = 4.65$ 

# BOD

## And we can use this for any statistic

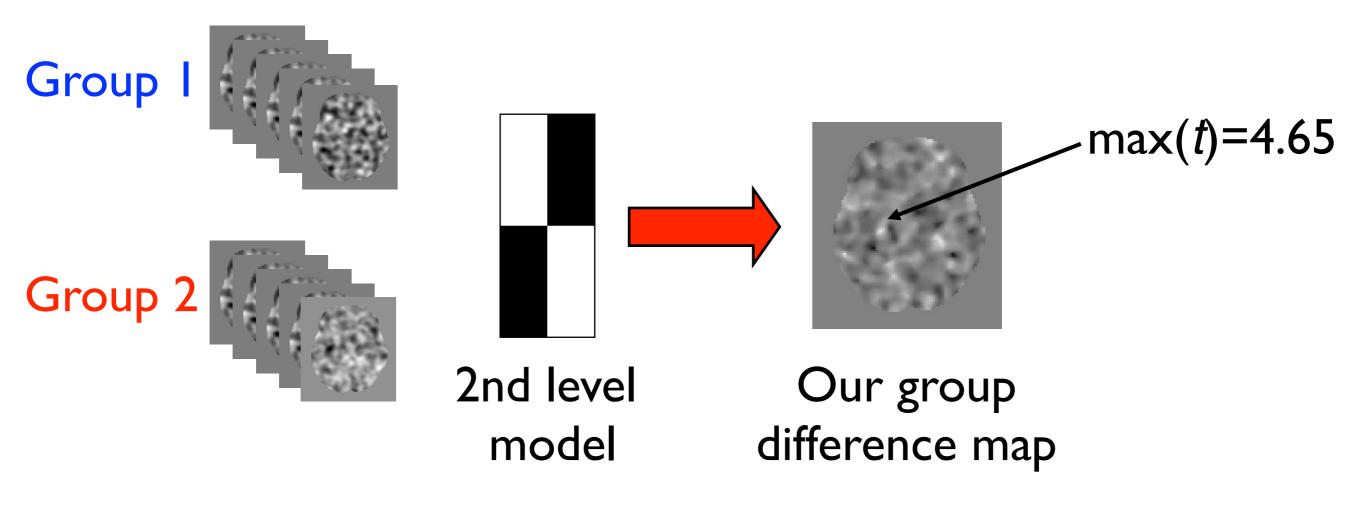
This is what we got

We compared activation by painful stimuli in two groups of 5 subjects each.



Very intriguing

activation.  $t_8 = 4.65$ 

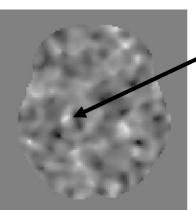


# BODE

## And we can use this for any statistic

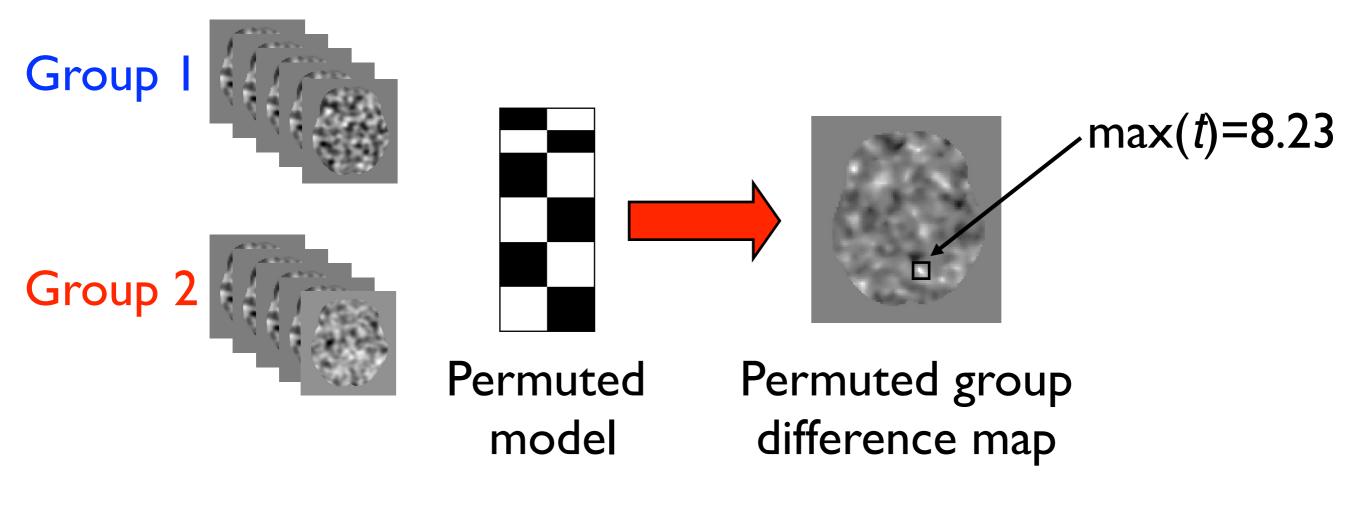
This is what we got

We compared activation by painful stimuli in two groups of 5 subjects each.



Very intriguing

activation.  $t_8 = 4.65$ 

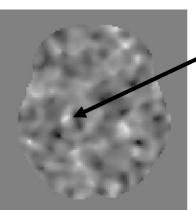


# BS

### And we can use this for any statistic

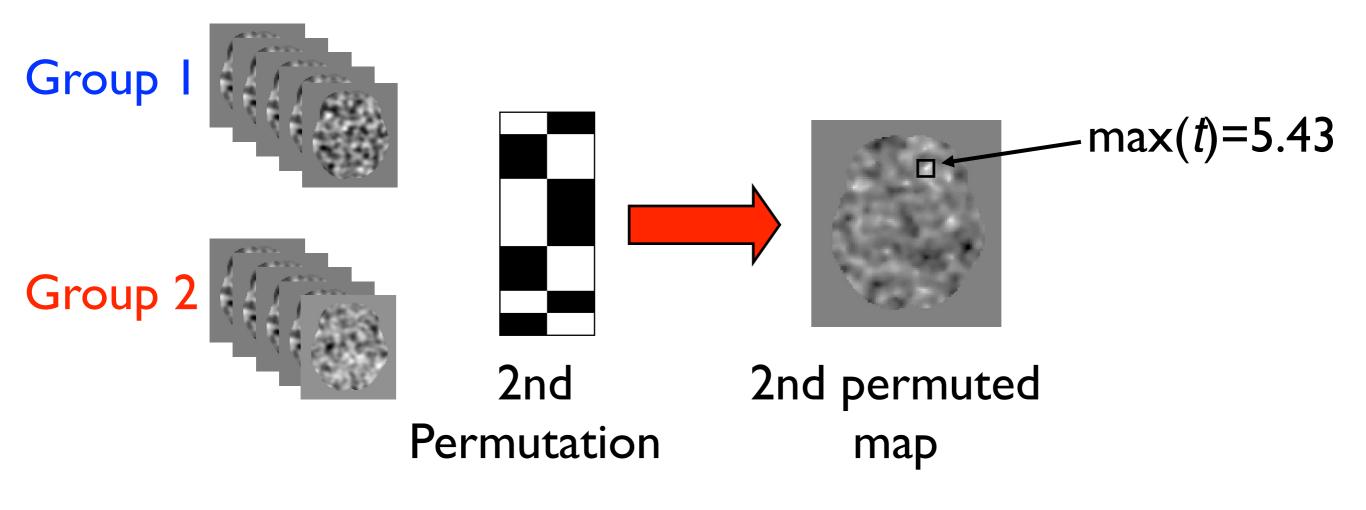
This is what we got

We compared activation by painful stimuli in two groups of 5 subjects each.



Very intriguing

activation.  $t_8 = 4.65$ 

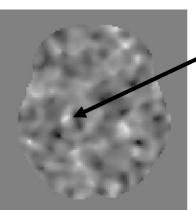


# ROL

#### And we can use this for any statistic

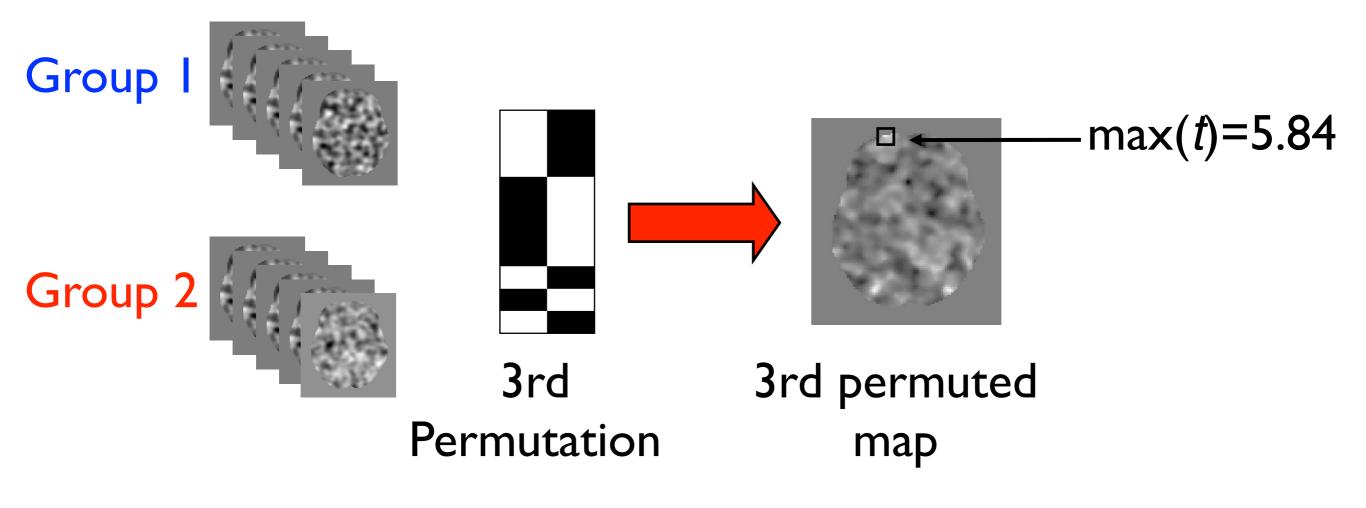
This is what we got

We compared activation by painful stimuli in two groups of 5 subjects each.



Very intriguing

activation.  $t_8 = 4.65$ 

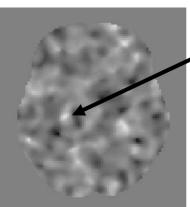


# BOD

## And we can use this for any statistic

This is what we got

We compared activation by painful stimuli in two groups of 5 subjects each.



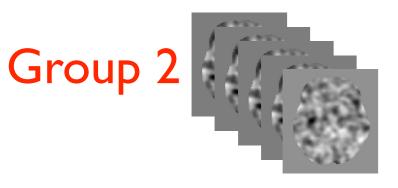
Original

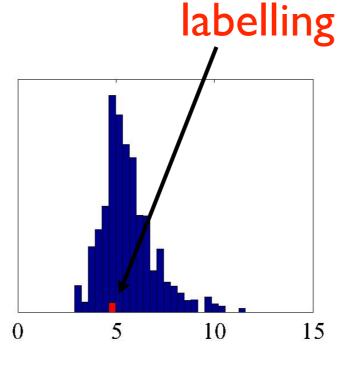
Very intriguing

activation.  $t_8 = 4.65$ 

Prof. ran to write to Nature Neuro. **But**, did they jump the gun?

Group I





5000 permutations

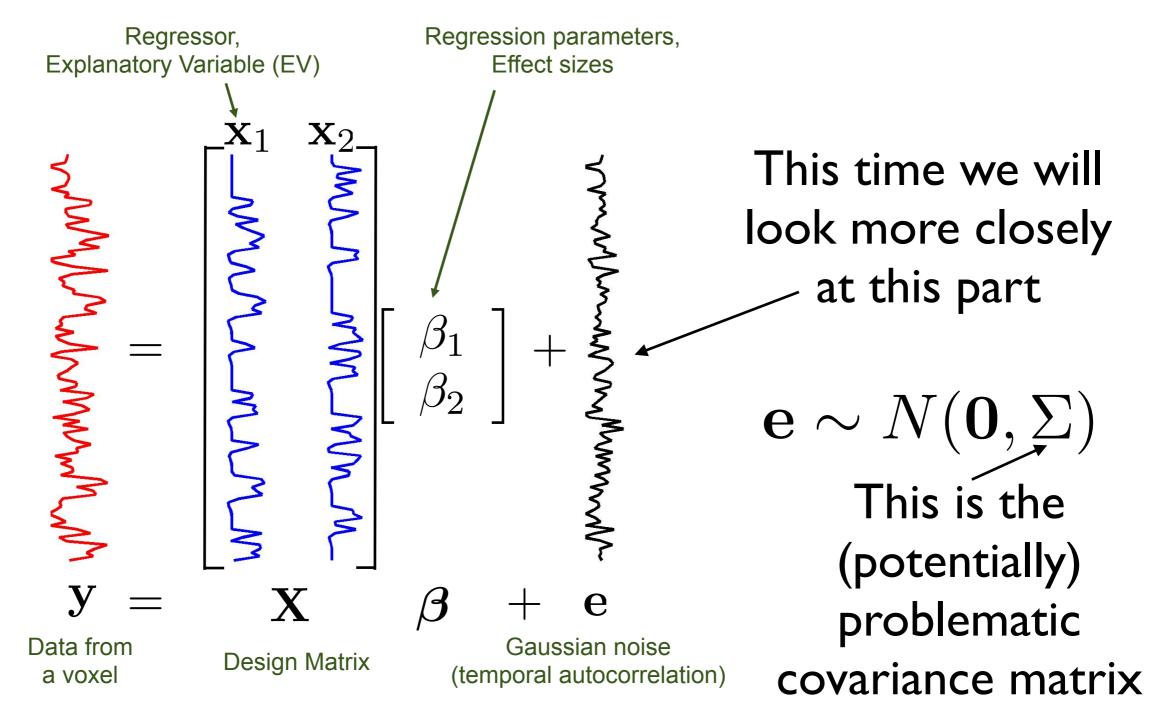
3925 permutations yielded higher max(t)-value than original labelling. We cannot reject the null-hypothesis.



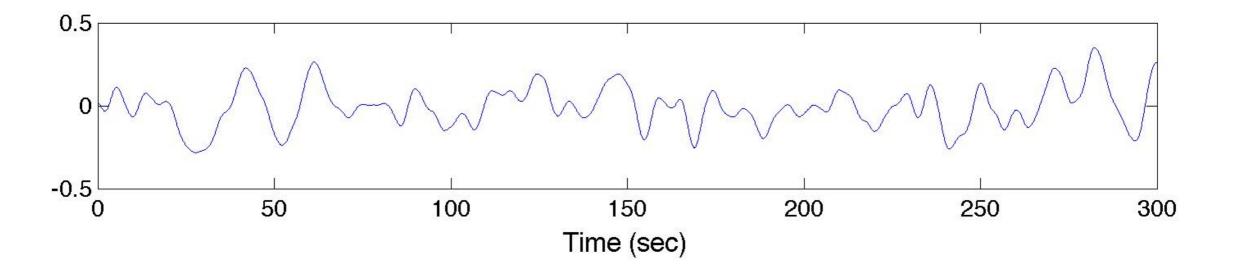
#### But beware the "exchangeability"

- When we swap the labels of two data-points we need to make sure that they are "exchangeable"
- "Exchangeable" means that the covariance matrix of the noise/error after model fitting isn't changed by a permutation (will show examples of this)

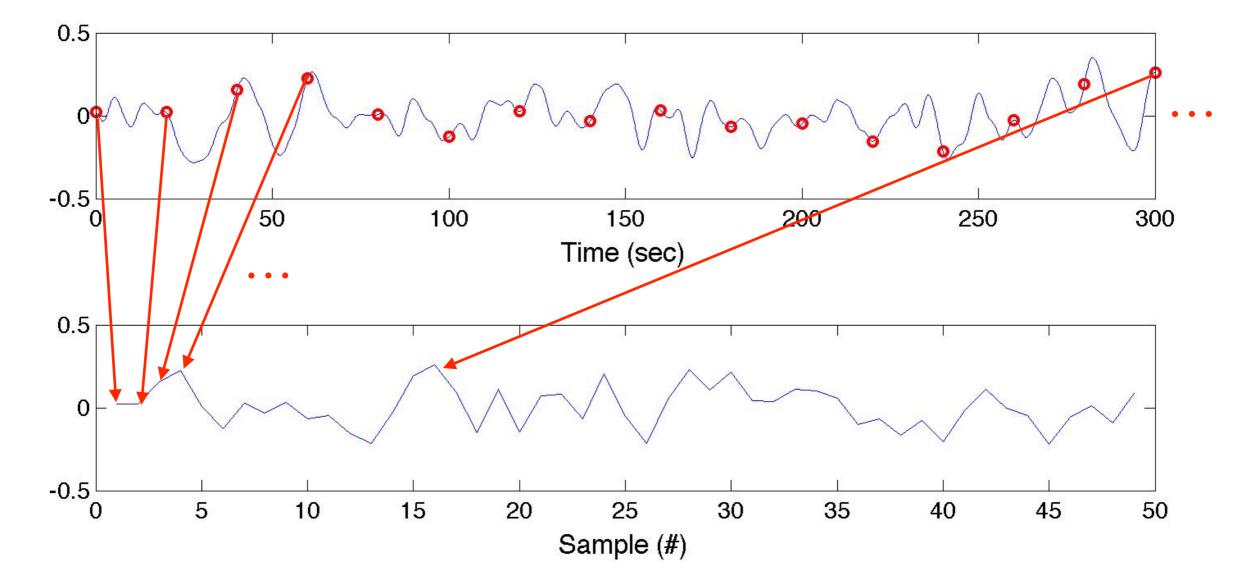
• You may, or may not, have seen this slide in the 1st level GLM talk.



• One important component of noise in fMRI consists of physiological/neuronal events convolved by the HRF

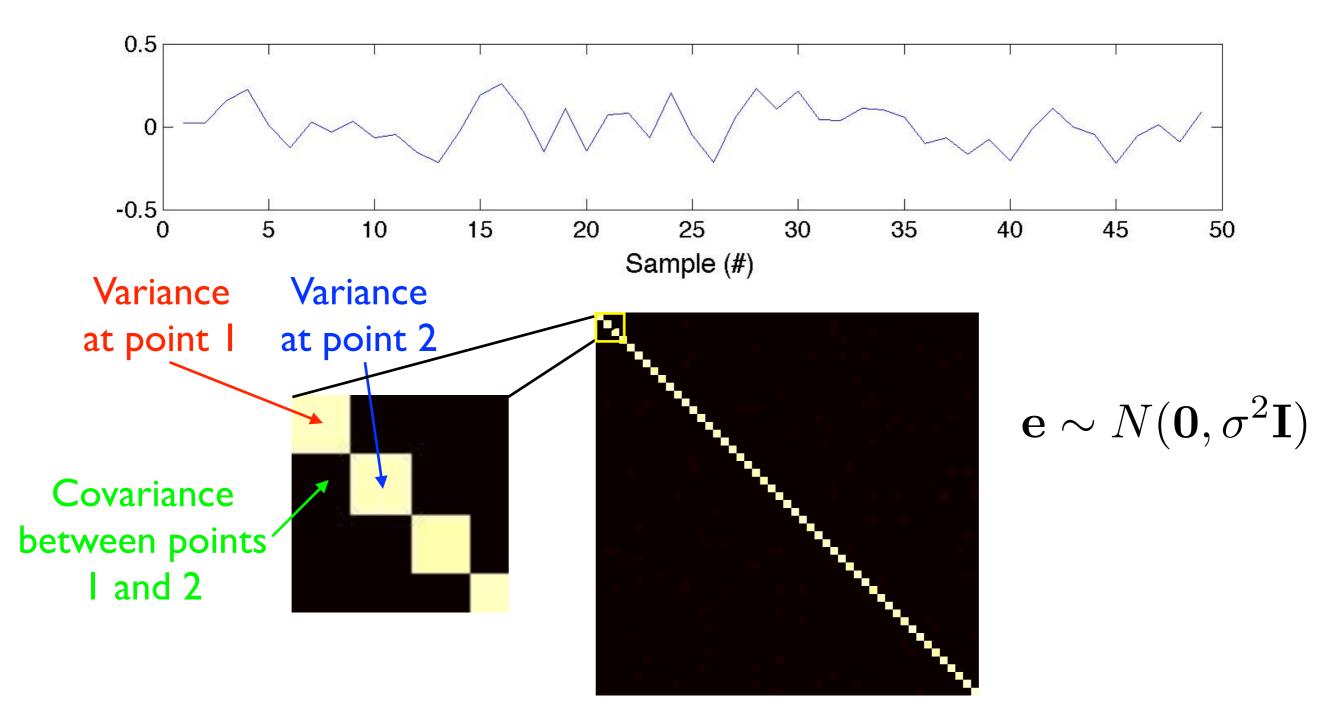


• One important component of noise in fMRI consists of physiological/neuronal events convolved by the HRF

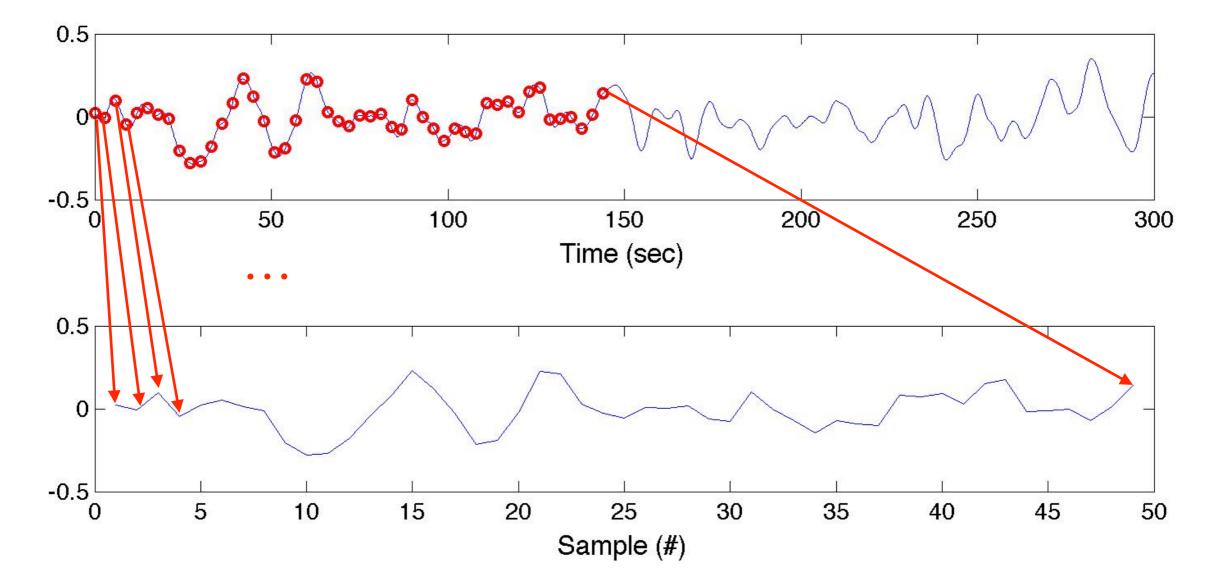


If we sample this every 20 seconds it no longer looks "smooth"

• One important component of noise in fMRI consists of physiological/neuronal events convolved by the HRF

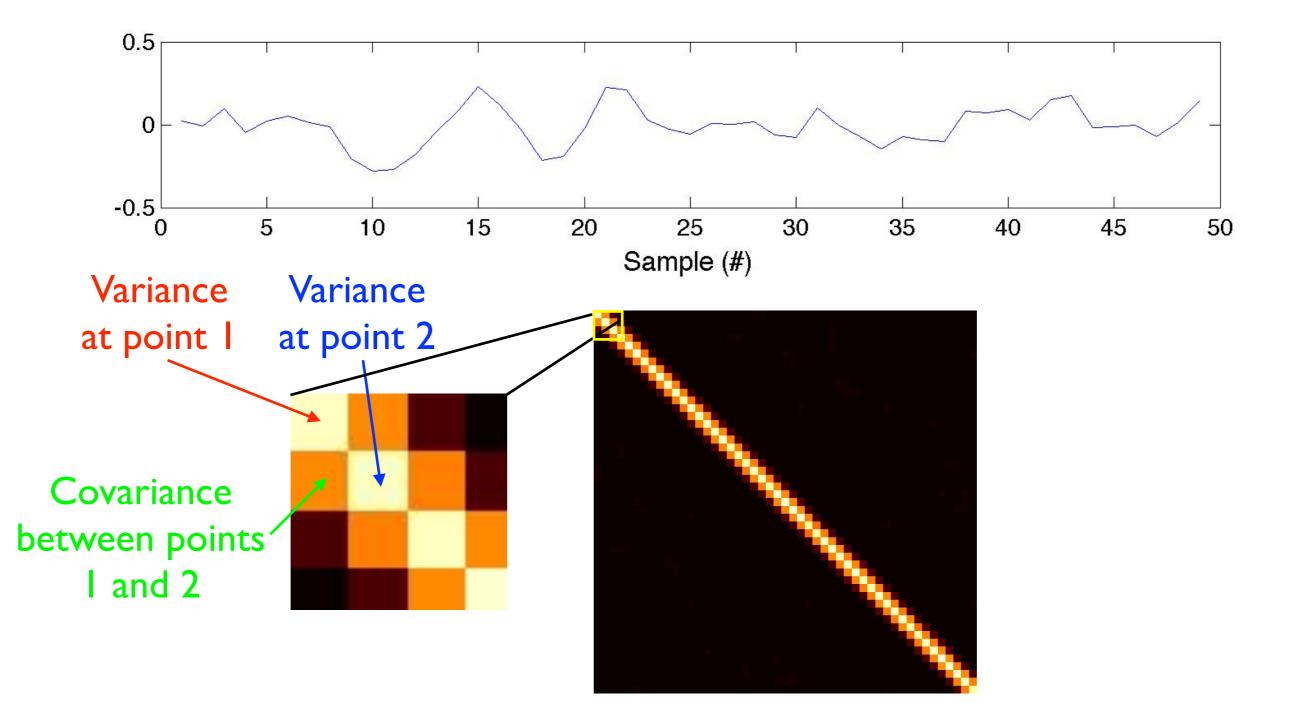


• One important component of noise in fMRI consists of physiological/neuronal events convolved by the HRF

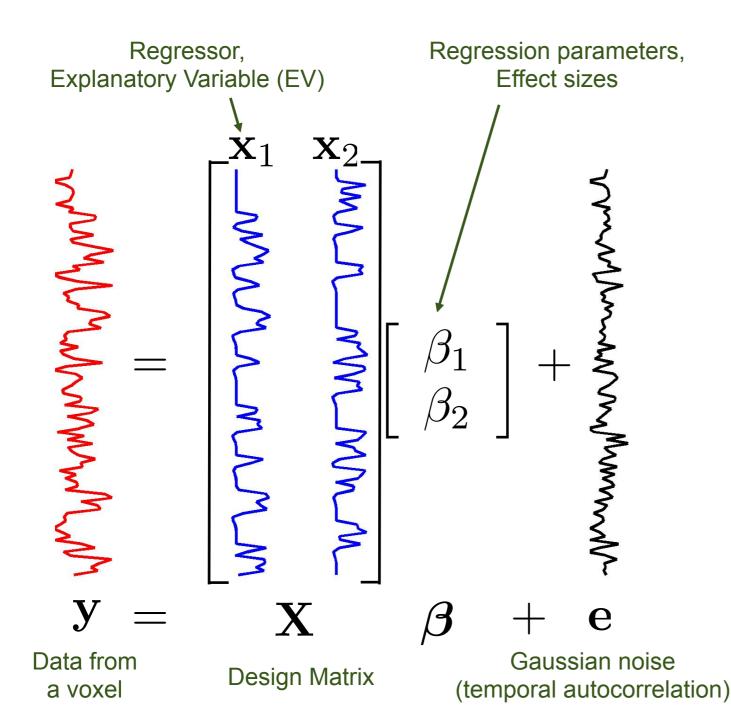


But that is not a realistic TR. What about every 3 seconds?

• One important component of noise in fMRI consists of physiological/neuronal events convolved by the HRF

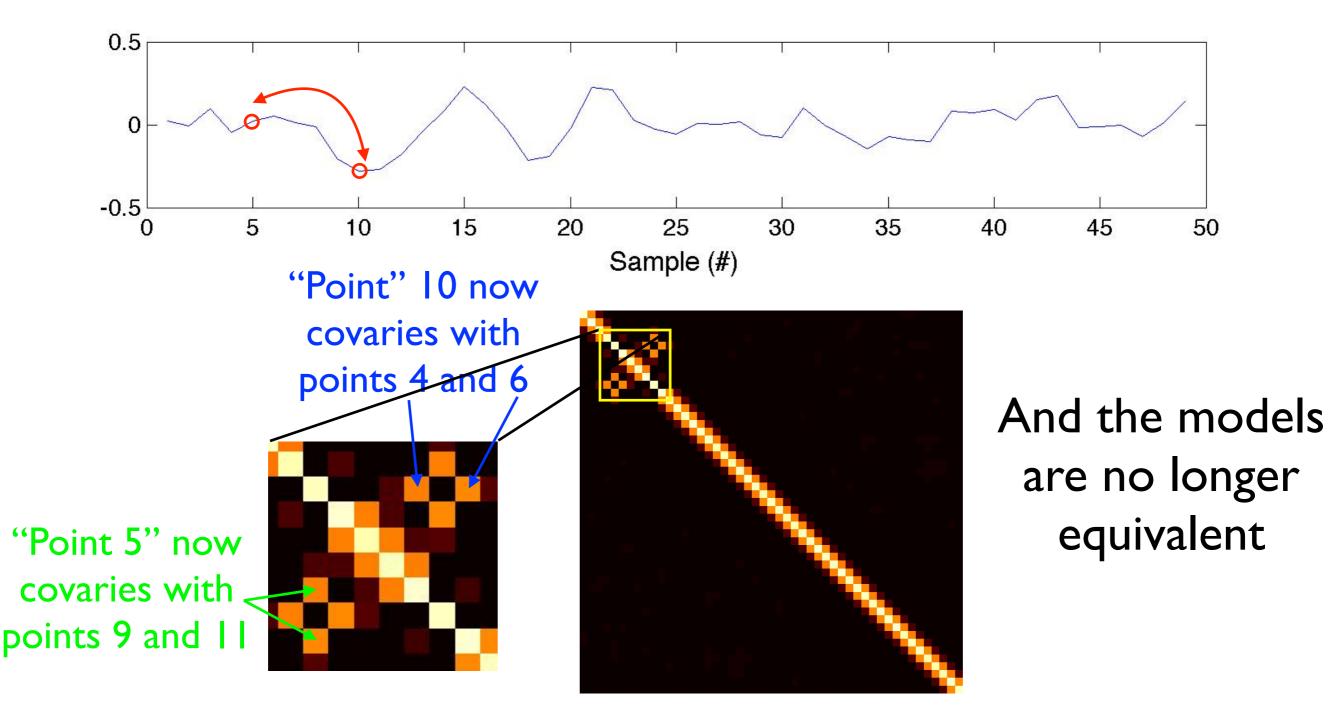


Let us now return to our model again

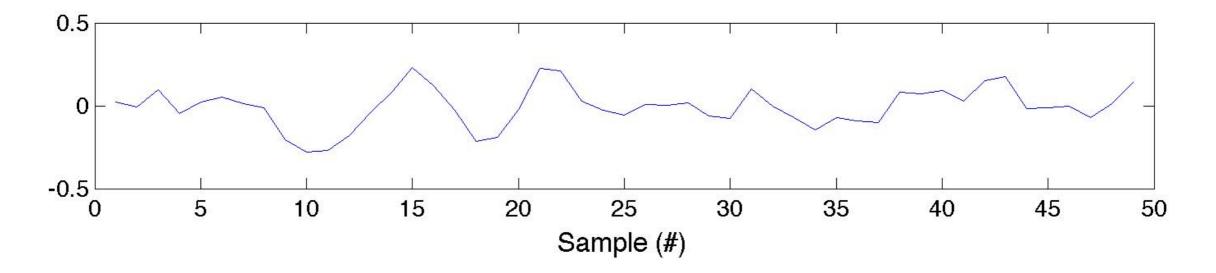


- The model consists of our regressors X and the noise model
- All permutations must result in "equivalent models"
- Let us now see what happens if we swap two data-points (points 5 and 10)

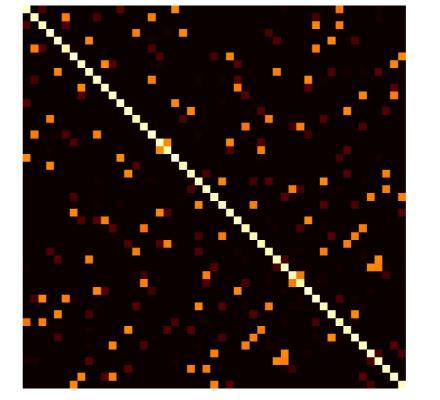
• One important component of noise in fMRI consists of physiological/neuronal events convolved by the HRF



• One important component of noise in fMRI consists of physiological/neuronal events convolved by the HRF



And for a random permutation ...



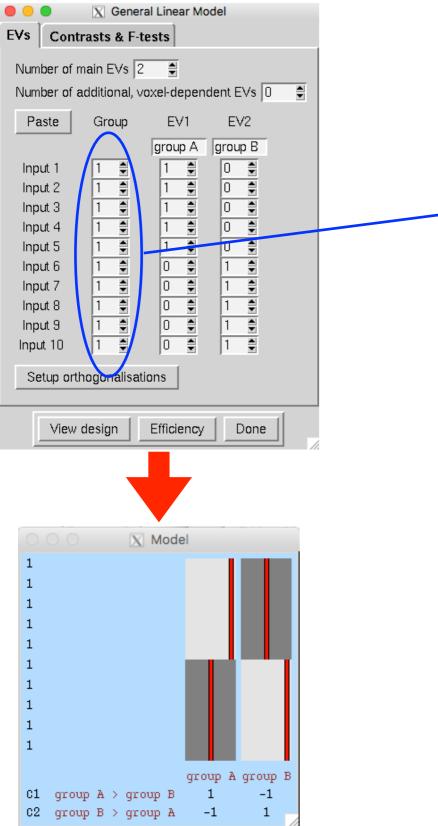
And the models are no longer equivalent



## Back to exchangeability

- Data-points are not "exchangeable" if swapping them means that the noise covariance-matrix ends up looking different.
- Formally "The joint distribution of the data must be unchanged by the permutations under the null-hypothesis".
- If the noise covariance-matrix has non-zero off-diagonal elements (covariances) you need to beware.
- You typically never estimate or see the covariancematrix. You need to "imagine it" and determine from that if there is a problem.

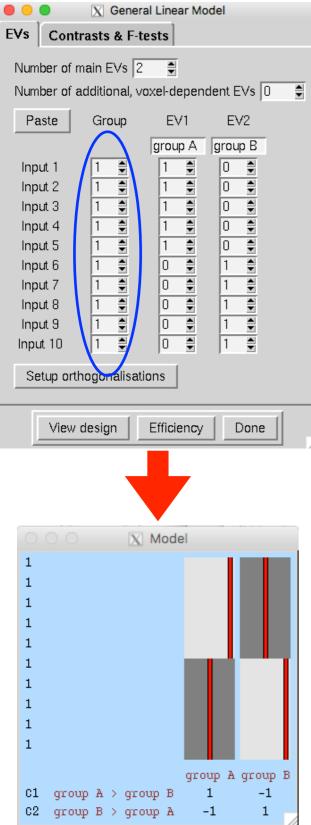




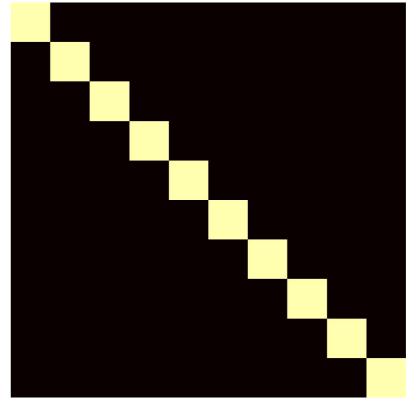
This is the "exchangeability group". Here all scans are in the same group, which means any scan can be exchanged for any other.

N.B. The "group" labelling is used for completely different purposes when using FLAME/GRFT



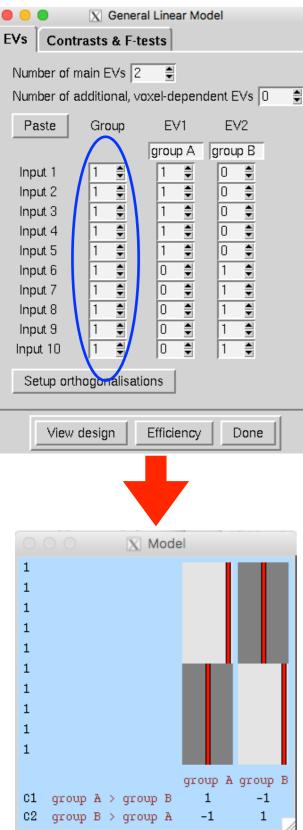


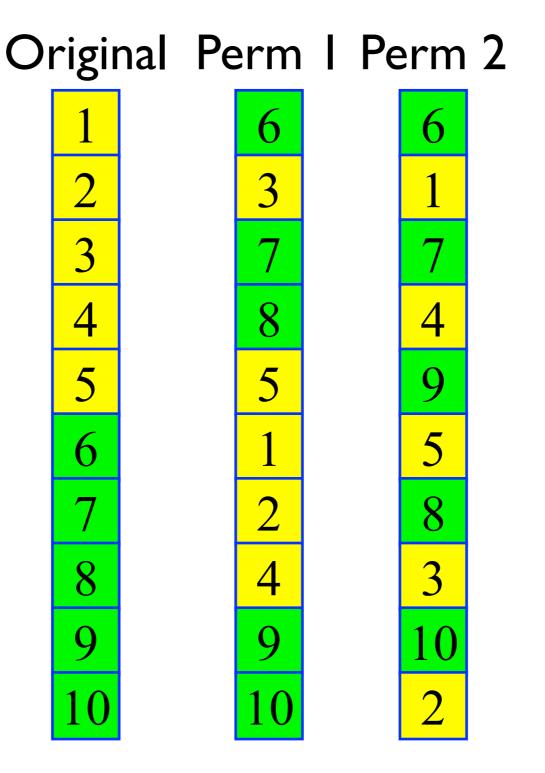
#### Assumed covariance matrix



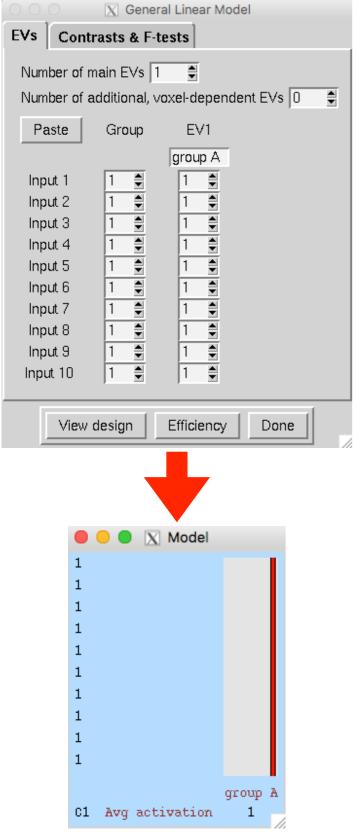
The implicit assumption here is that data from all subjects have the same uncertainty and are all independent







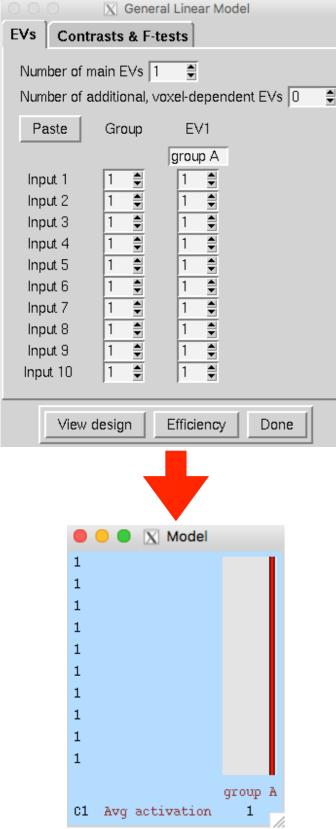


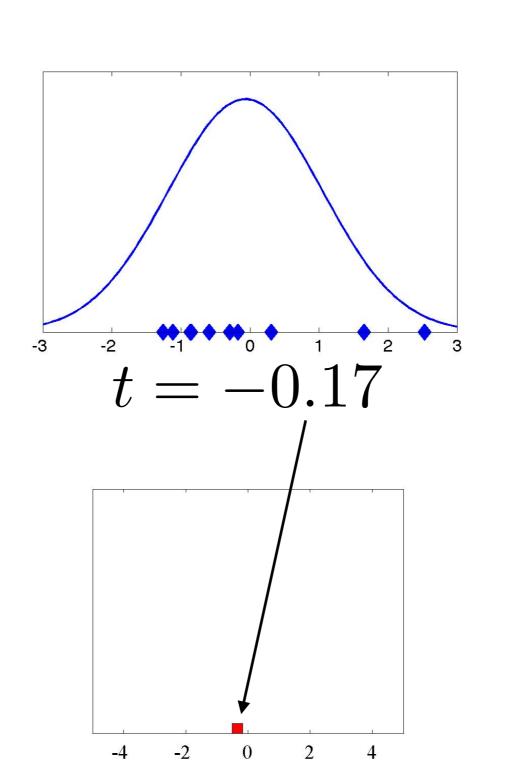


Here we model a single mean and want to know if that is different from zero

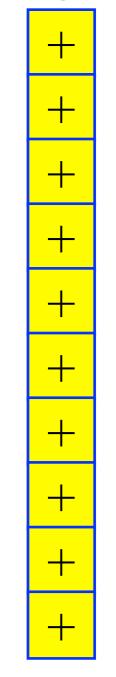
But there isn't really anything to permute, or is there?



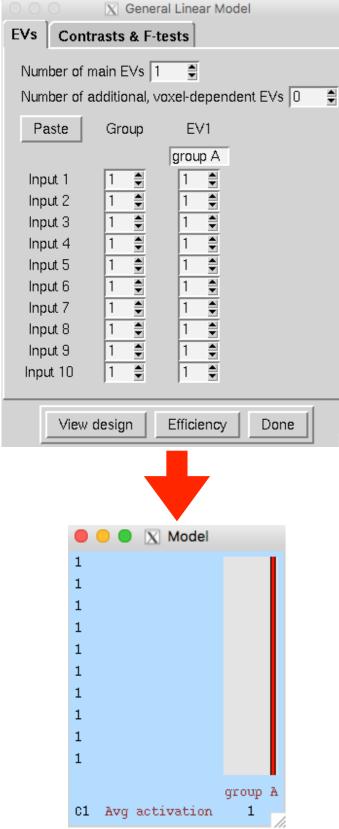


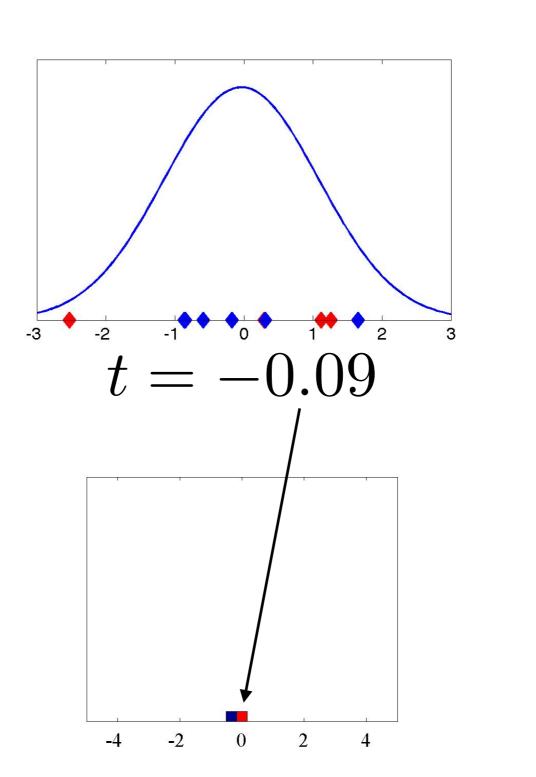


Original

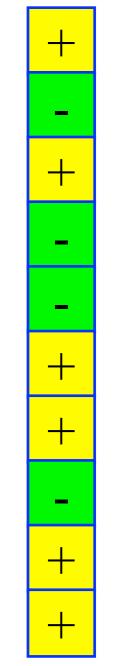




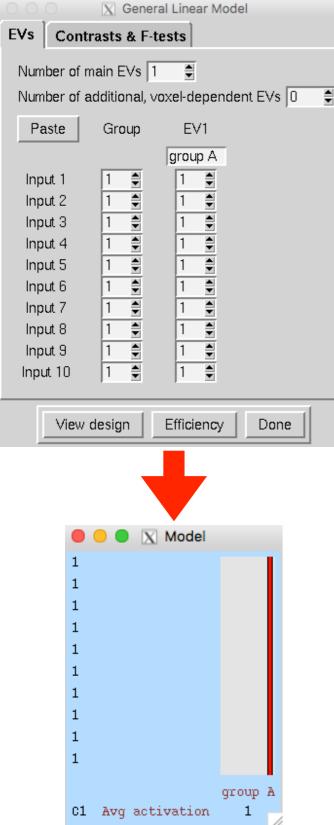


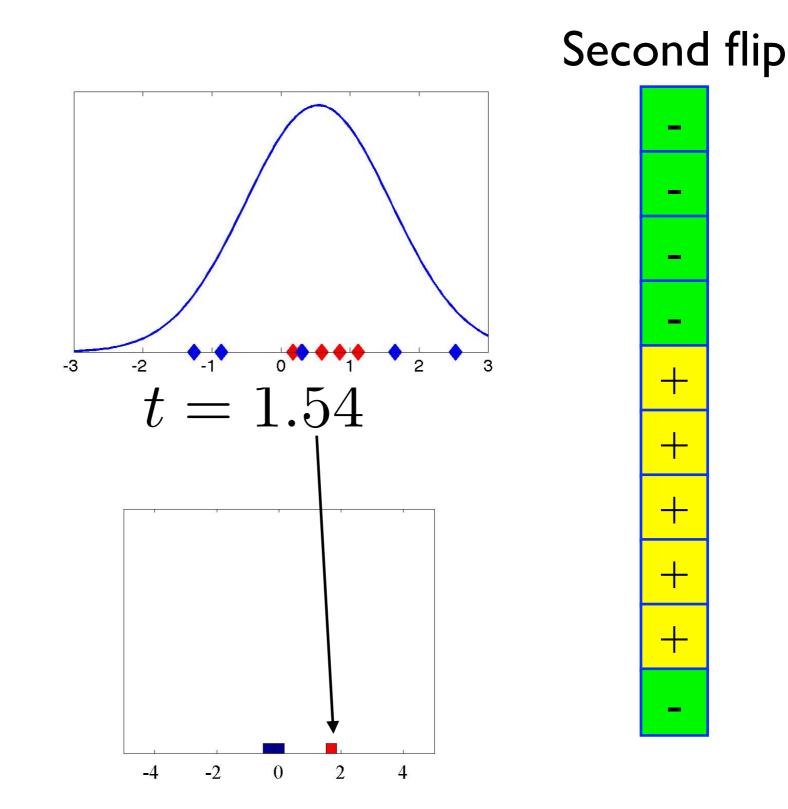


First flip

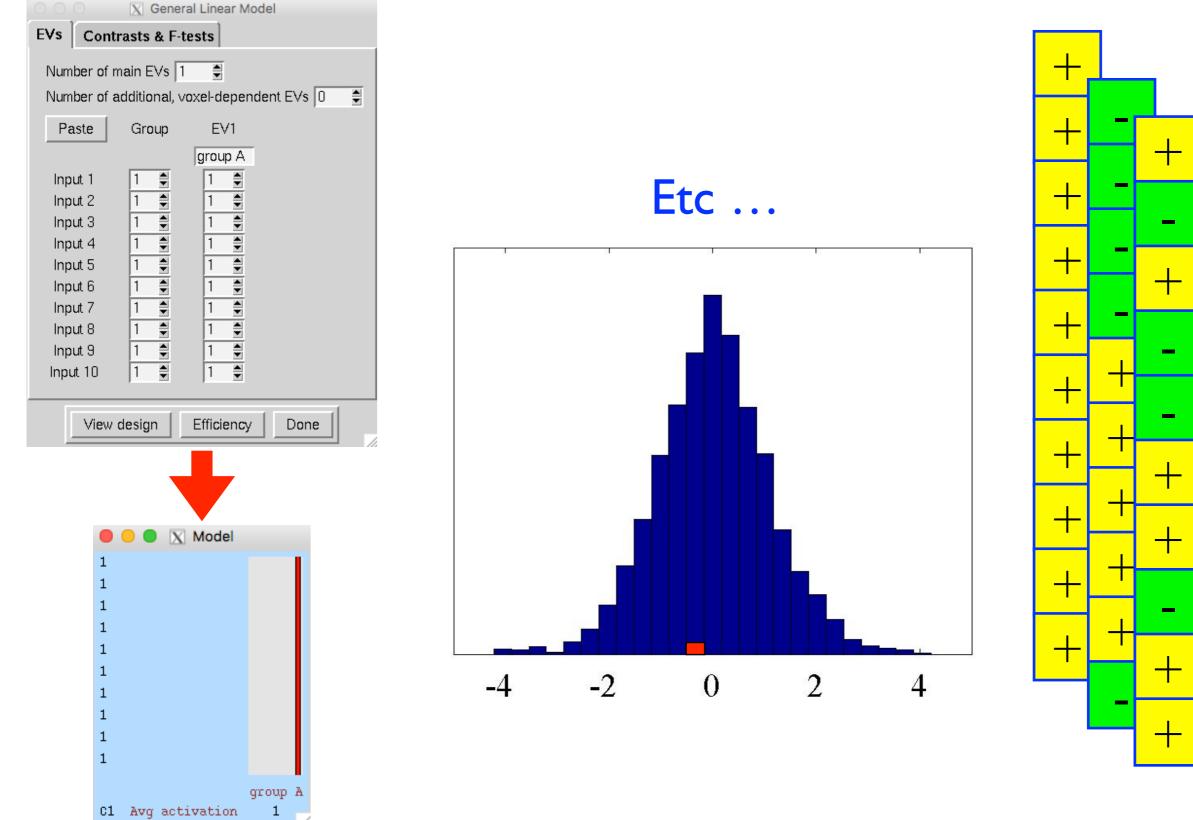




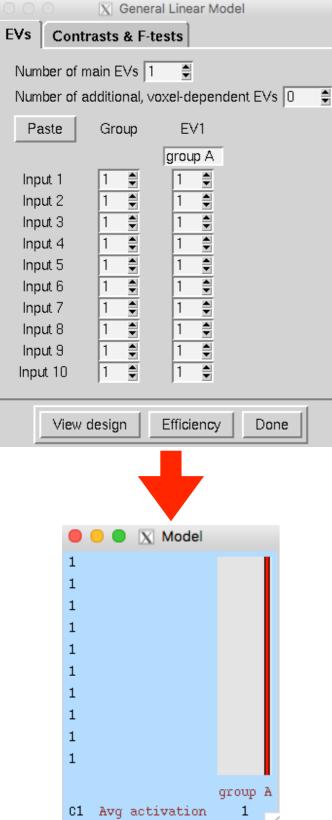


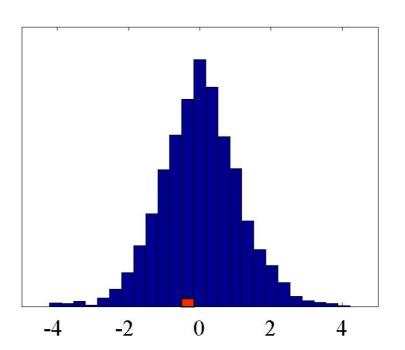












And the assumptions are:

- Symmetric errors
- Errors independent
- Subjects drawn from a single population



8 😑 🖶	X General Line	ar Model	
EVs Contrasts & F-tes	ts		
Number of main EVs 6			
Number of additional, vox		<b></b>	
	. ,		
Paste Group	EV1 EV2	EV3 EV4	EV5 EV6
			Subj 4 Subj 5
Input 1       1         Input 2       1         Input 3       2         Input 4       2         Input 5       3         Input 6       3         Input 7       4         Input 8       4         Input 9       5			
Input 3 2			
Input 4 2			
Input 5 3 🛢	1 🛔 🛛 🗘		0
Input 6 🛛 3 🌻	-1 🛊 🛛 🌲	0 🛊 1 🌻	0 🛊 0 🌲
Input 7 4 🚔			
Input 8 4 🖨 Input 9 5 🖨	-1 ♣ 0 ♣ 1 ♣ 0 ♣		
Input 9 5 🚔 Input 10 5 🚔			
	· - · - ·		
Setup orthogonalisation	15		
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Here we can only exchange scans within each subject. I.e. Input 1 for Input 2, Input 3 for Input 4 etc



8 - 🖶				X Ger	eral Line	ar Mode	1				
EVs C	ontra	ists & F-	tests								
Number	of ma	ain EVs [	6								
				ependent	EVs 0	<b></b>					
Paste	:	Group	E∖	/1 E	V2	EV3	EV4	I	EV5	EV6	
Input 1 Input 2 Input 4 Input 4 Input 5 Input 6 Input 9 Input 1 Setup	2 3 4 5 7 3 3 0	1 1 1 1 2 1 2 1 3 1 3 1 4 1 5 1 5 1 5 1 5 1 5 1 5 1 5 1 5	A>B 1 -1 1 -1 1 -1 1 -1 1 -1 1 -1 1 -1 1 -1 1 -1 -			Subj 2       0       1       1       1       0       1       0       1       0	Subj 3	Su 0 0 0 1 1 0 0		Subj 5	_
			View	design	Effici	ency	Done				
					X Mod	el					
	1 2 3 3 4 4 5 5		A>B	Subj 1							

Assumed covariance matrix



The implicit assumption here is that data from all subjects have the same uncertainty and that there is no dependence between subjects



8 😑 🖶			X Ger	neral Linea	r Model				
EVs C	ontrasts &	F-tests							
Number	of main EV	s 6 🌻							
	of additiona			EVs 0					
Paste				,	EV3	EV4	EV5	EV6	
1 4316		A>B	Sub			Subj 3	Subj 4	Subj 5	
Input 1	1	- · ·							
Input 2		-1							
Input 3	2	1	0	1	<b></b>	0	0 🛢	0	
Input 4	1 2	-1	0	1	-	0 🛢	0 🛢	0 🚔	
Input 5	3	1	<u> </u>			1	0	0	
Input 6	i 3 4	-1				1	0 🚔 1 🚔		
Input 7 Input 8	3 4	-1					1		
Input 9	5		■ 0				0	1	
Input 1	0 5	-1	0			0	0	1	
Setup orthogonalisations									
View design Efficiency Done									
		View	/ design	Efficie	ncy	Done			
		View	/ design	Efficie	ncy	Done		7	
		View	/ design	Efficie	ncy	Done		1	
		View	/ design	Efficie	ncy	Done			
		View		Efficie	•	Done			
	•••	View		♥	•	Done			
	<ul> <li>•••••</li> <li>•••••</li> <li>1</li> <li>1</li> </ul>	View		♥	•	Done			
	1 2	View		♥	•	Done			
	1 2 2	View		♥	•	Done			
	1 2 2 3	View		♥	•	Done			
	1 2 2 3 3	View		♥	•	Done			
	1 2 2 3 3 4	View		♥	•	Done			
	1 2 2 3 3	View		♥	•	Done			
	1 2 3 3 4 4			♥	•	Done			
	1 2 3 3 4 4 5			X Mode					
	1 2 3 3 4 4 5	View A>B 1		♥			4 Subj 5	5	

Assumed covariance matrix

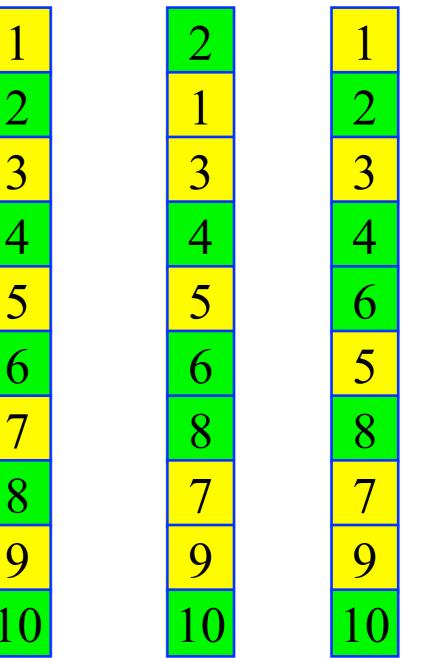


The implicit assumption here is that data from all subjects have the same uncertainty and that there is no dependence between subjects



8 - 3				X Ge	eneral Lir	near Mo	del					
EVs Contrasts & F-tests												
Number	of ma	ain EVs	6 🚔									
Number	of ad	Iditional	, voxel-de	penden	t EVs 🛛							
Paste	:	Group	EV	1	EV2	EV3	I	EV4	E∖	/5	ΕV	6
Input 1 Input 2 Input 3 Input 4 Input 5 Input 6 Input 9 Input 9 Input 1	2 3 4 5 7 3 3 0	1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1	A>B 1 -1 1 -1 1 -1 1 -1 1 -1 1 -1 -			Subj 2			Subj 0 0 0 0 1 1 0 0		Subj 5 0 0 0 0 0 0 0 1 1	
View design Efficiency Done												
	•				X Mo	del						
	1 2 3 4 4 5 5								I			
	C1	A>B	A>B 1	Subj 1	0	2 Sul	Ō	0	-1 30	ubj9 0	, 	
	C2	B>A	-1	0	0		0	0		0	1	

Original Perm I Perm 2 ···





# Outline

- Null-hypothesis and Null-distribution
- Multiple comparisons and Family-wise error
- Different ways of being surprised
  - Voxel-wise inference (Maximum z)
  - Cluster-wise inference (Maximum size)
- Parametric vs non-parametric tests
- Enhanced clusters
- FDR False Discovery Rate

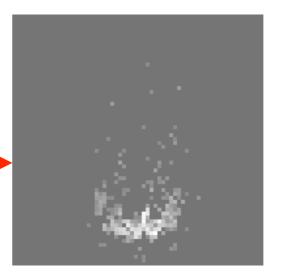


# Clustering cookbook

Instead of resel-based correction, we can do clustering:



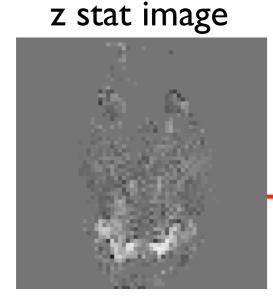
Threshold at (arbitrary!) z level



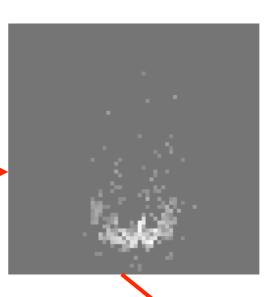


# Clustering cookbook

Instead of resel-based correction, we can do clustering



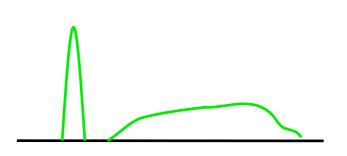
Threshold at (arbitrary!) z level



Form clusters from surviving voxels. Calculate the size threshold u(R,z). Any cluster larger than u "survives" and we reject the null-hypothesis for that.



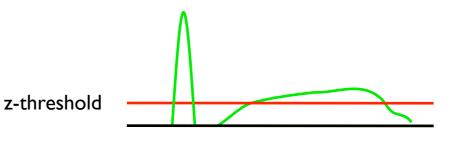
This is arbitrary and a trade-off





This is arbitrary and a trade-off

I. Low threshold - can violate RFT assumptions, but can detect clusters with large spatial extent and low z

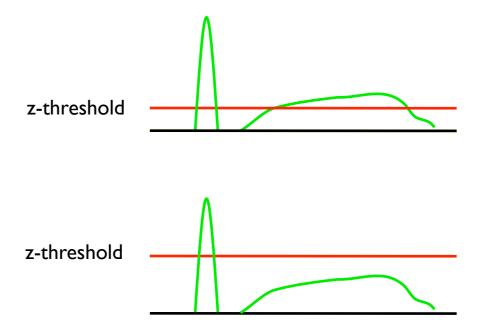




This is arbitrary and a trade-off

 Low threshold - can violate RFT assumptions, but can detect clusters with large spatial extent and low z

2. **High threshold** - gives more power to clusters with small spatial extent and high z





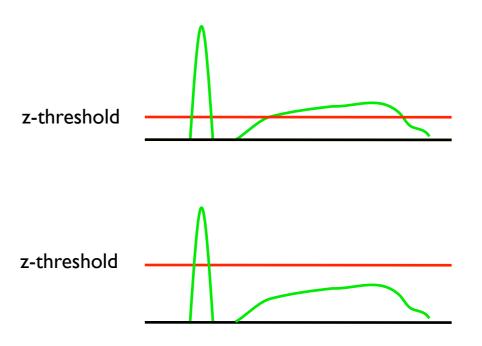
This is arbitrary and a trade-off

I. Low threshold - can violate RFT assumptions, but can detect clusters with large spatial extent and low z

2. **High threshold** - gives more power to clusters with small spatial extent and high z

Tends to be more sensitive than voxel-wise corrected testing

Results depend on extent of spatial smoothing in pre-processing





### TFCE

#### **Threshold-Free Cluster Enhancement**

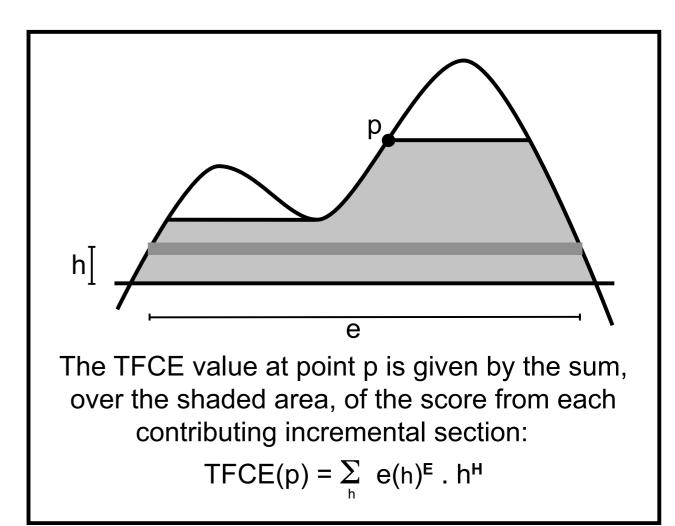
[Smith & Nichols, NeuroImage 2009]

- Cluster thresholding:
  - popular because it's sensitive, due to its use of spatial extent
  - but the pre-smoothing extent is arbitrary
  - and so is the cluster-forming threshold

unstable and arbitrary

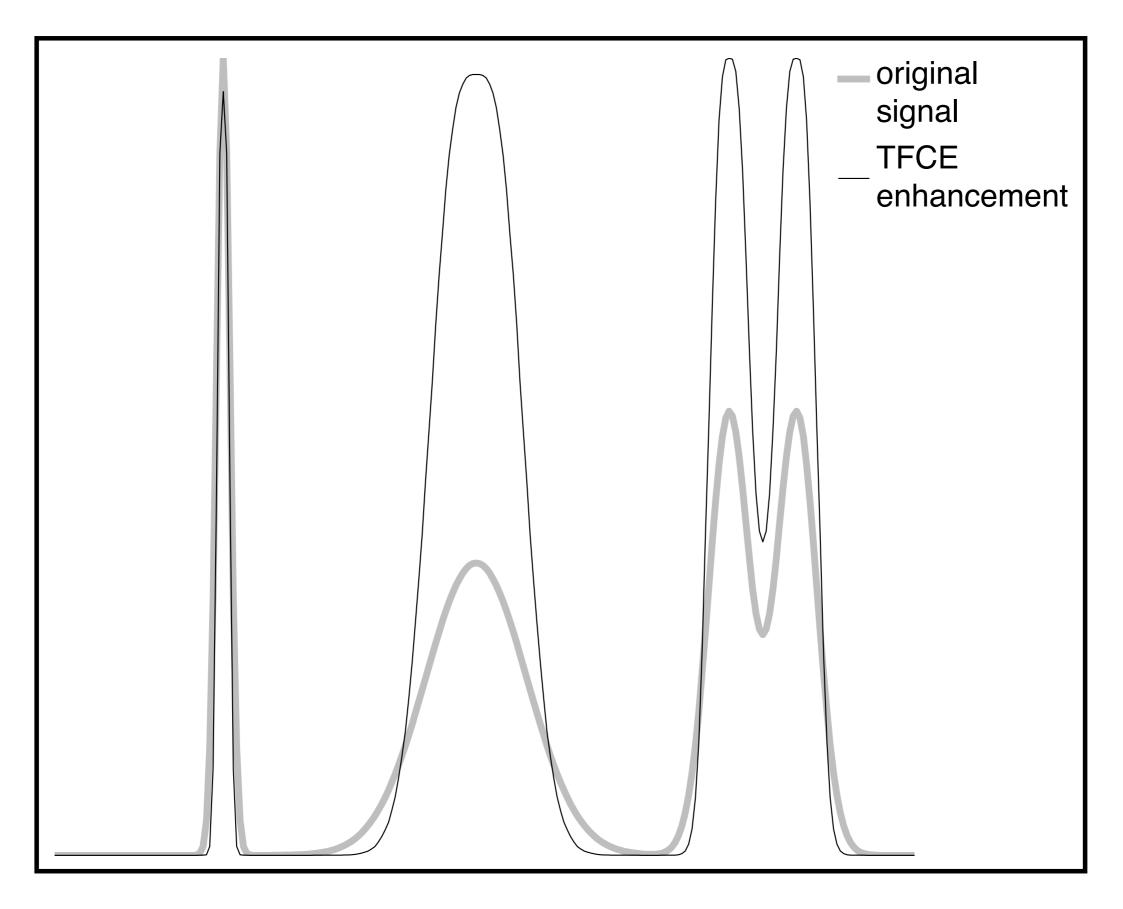
### • TFCE

- integrates cluster "scores" over all possible thresholds
- output at each voxel is measure of local cluster-like support
- similar sensitivity to optimal cluster-thresholding, but stable and non-arbitrary



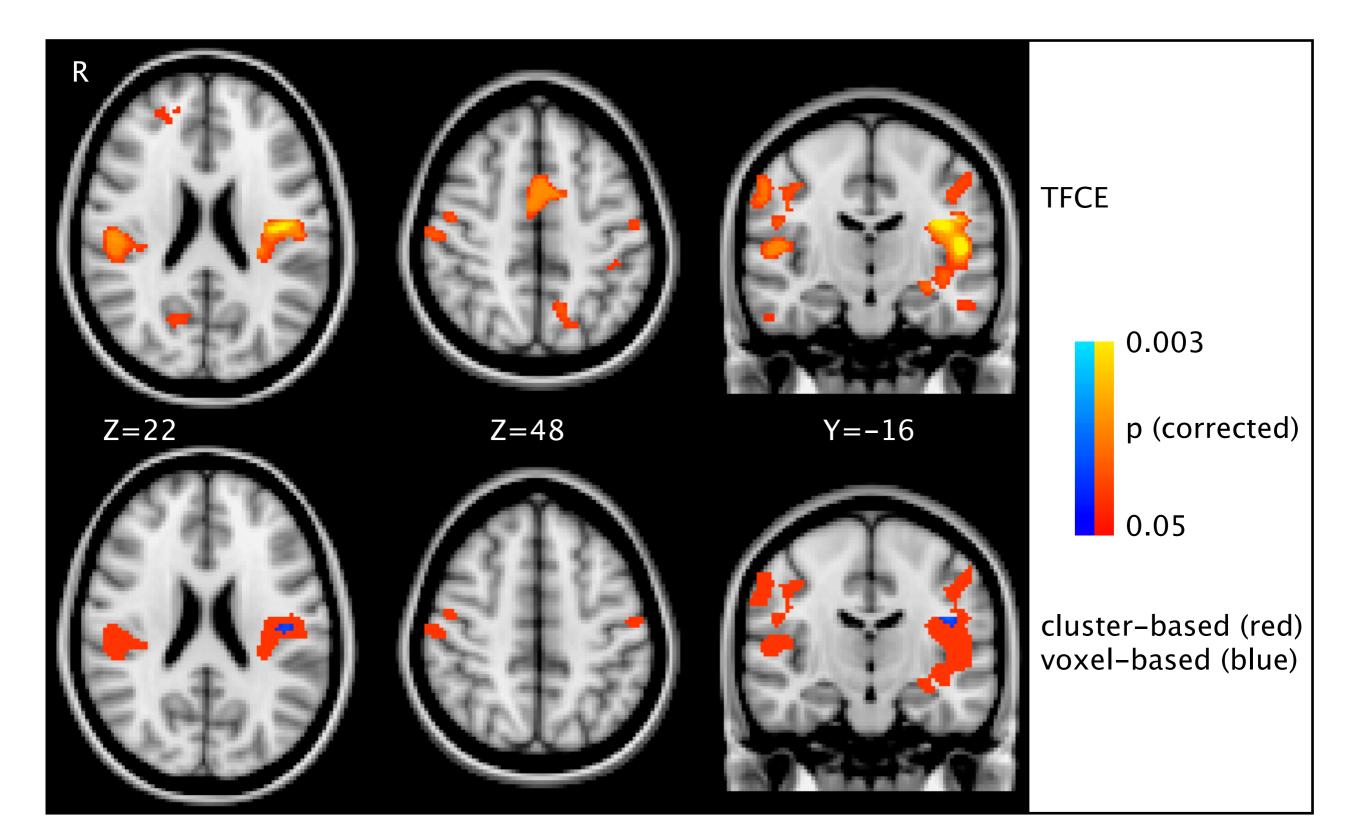


#### Qualitative example





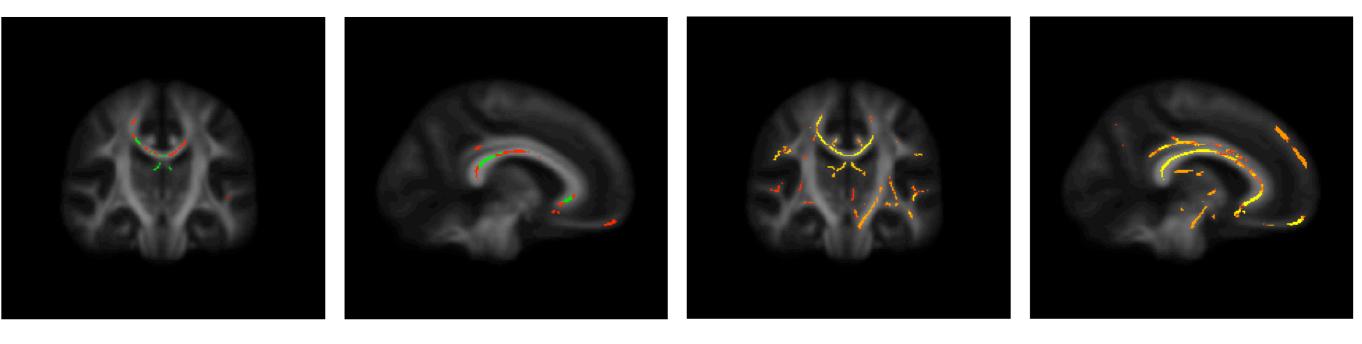
### TFCE for FSL-VBM

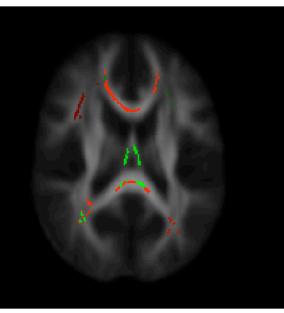




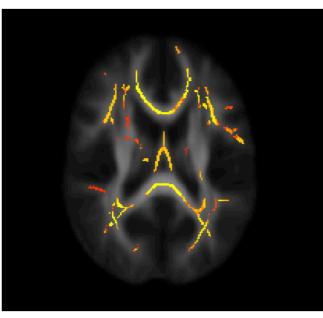
### TFCE for TBSS

### controls > schizophrenics p<0.05 corrected for multiple comparisons across space, using randomise





cluster-based: cluster-forming threshold = 2 or 3



#### TFCE



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- FDR False Discovery Rate



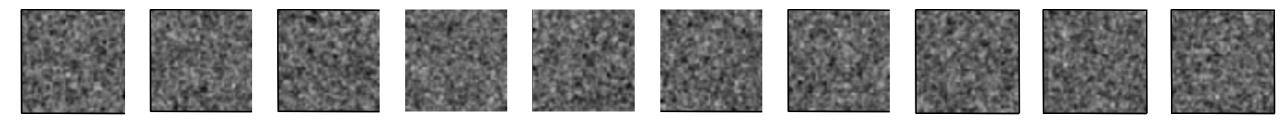
### False Discovery Rate



- FDR: False Discovery Rate A "new" way to look at inference.
- Uncorrected (for multiple-comparisons):
  - Is equivalent to saying: "I am happy to nearly always say something silly about my experiments".
  - On average, **5% of all voxels** are false positives
- Family-Wise Error (FWE):
  - Is equivalent to saying: "I am happy to say something silly about 5% of my experiments".
  - On average, **5% of all experiments** have one or more false positive voxels
- False Discovery Rate
  - Is equivalent to saying: "I am happy if 5% of what I say about each experiment is silly".
  - On average, **5% of significant voxels** are false positives

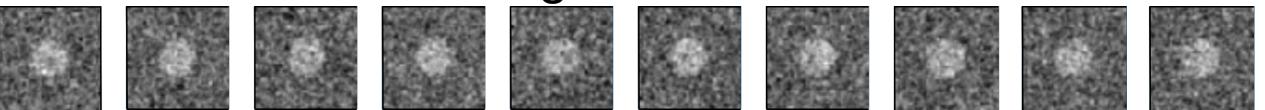
## Little imaging demonstration.

#### Noise



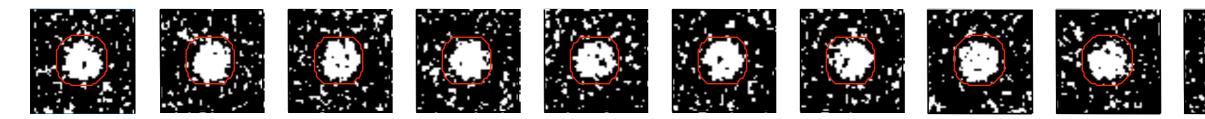
## Signal

#### Signal+Noise





#### uncorrected voxelwise control of FP rate at 10%



percentage of all null pixels that are False Positives

### control of FamilyWise Error rate at 10%



occurrence of FamilyWise Error

FWE

### control of False Discovery Rate at 10%



percentage of activated (reported) pixels that are False Positives



### FDR for dummies

- Makes assumptions about how errors are distributed (like GRT).
- Used to calculate a threshold.
- Threshold such that X% of super-threshold (reported) <u>voxels</u> are false positives.
- Threshold depends on the data. May for example be very different for [1 0] and [0 1] in the same study.